

9/5/2013

TODAY: PARAMETER II

I: MAP / BAYESIAN ESTIMATION

II: MIXTURE DENSITIES

III: VARIATIONAL BAYES & EM

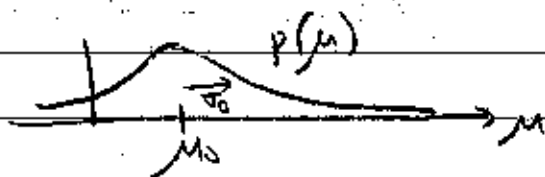
ML ESTIMATION: $\hat{\theta} = \operatorname{argmax} p(\mathcal{X}|\theta)$ MAP ESTIMATION $\hat{\theta} = \operatorname{argmax} p(\theta|\mathcal{X})$
 $= \operatorname{argmax} \frac{p(\mathcal{X}|\theta)p(\theta)}{p(\mathcal{X})}$

$$= \operatorname{argmax} p(\mathcal{X}|\theta) p(\theta)$$

$$= \operatorname{argmax} (\ln p(\theta) + \sum_{n=1}^N \ln p(x_n|\theta))$$

WHAT IS $p(\theta)$?

- ① CENTROID = VALUE OF θ YOU CONSIDER MOST PROBABLE



- ② SPREAD $\propto \frac{1}{\text{CONFIDENCE}}$ OF YOUR A PRIORI ESTIMATE

- ③ CONJUGATE PRIOR $\equiv p(\theta)$ CHOSEN SO THAT $\nabla_{\theta} \ln p(\theta)$ HAS THE SAME DEPENDENCE ON θ AS $\nabla_{\theta} \ln p(x_n|\theta)$

EXAMPLES

① GAUSSIAN MEAN

CONJUGATE PRIOR = GAUSSIAN

$$p(\mu) = \frac{1}{(2\pi)^{D/2} |\Sigma_0|^{1/2}} e^{-\frac{1}{2} (\mu - \mu_0)^T \Sigma_0^{-1} (\mu - \mu_0)}$$

$\mu_0, \Sigma_0 =$

$$\nabla_{\mu} (\ln p(\mu) + \sum_{n=1}^N \ln p(\vec{x}_n | \mu))$$

"HYPERPARAMETERS"

$$= \Sigma_0^{-1} (\mu_0 - \mu) + \Sigma^{-1} \sum_{n=1}^N (\vec{x}_n - \mu)$$

$$= 0 \quad \text{IF} \quad \boxed{\begin{matrix} \hat{\mu} = (\mathbf{I} + N)^{-1} (\mathbf{I} \mu_0 + \sum_{n=1}^N \vec{x}_n) \\ \mathbf{T} = \sum \Sigma^{-1} \end{matrix}}$$

② GAUSSIAN COVARIANCE

CONJUGATE PRIOR = WISHART DISTRIBUTION

$$p(\Sigma) \propto \frac{|R_0|^{D/2}}{|\Sigma|^{D/2}} e^{-\frac{\alpha}{2} \text{trace}(\Sigma^{-1} R_0)}$$

HYPERPARAMETERS
 α, R_0

$$p(\vec{x}_n | \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} \text{trace}((\vec{x}_n - \mu)^T \Sigma^{-1} (\vec{x}_n - \mu))}$$

$$= \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} \text{trace}(\Sigma^{-1} (\vec{x}_n - \mu) (\vec{x}_n - \mu)^T)}$$

TRACE IDENTITY: $\text{trace}(AB) = \text{trace}(BA)$

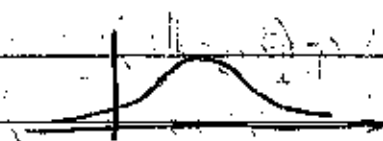
$$\therefore \hat{\Sigma} = \frac{1}{\alpha + N} \left(\alpha R_0 + \sum_{n=1}^N (\vec{x}_n - \hat{\mu}) (\vec{x}_n - \hat{\mu})^T \right)$$

CONFIDENCE

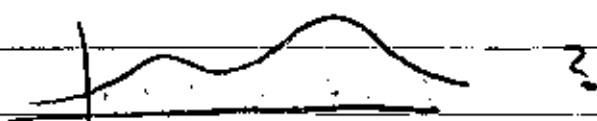
PRIOR BELIEF

II. MIXTURE DENSITIES

GAUSSIAN, LAPLACIAN, ... → UNIMODAL!

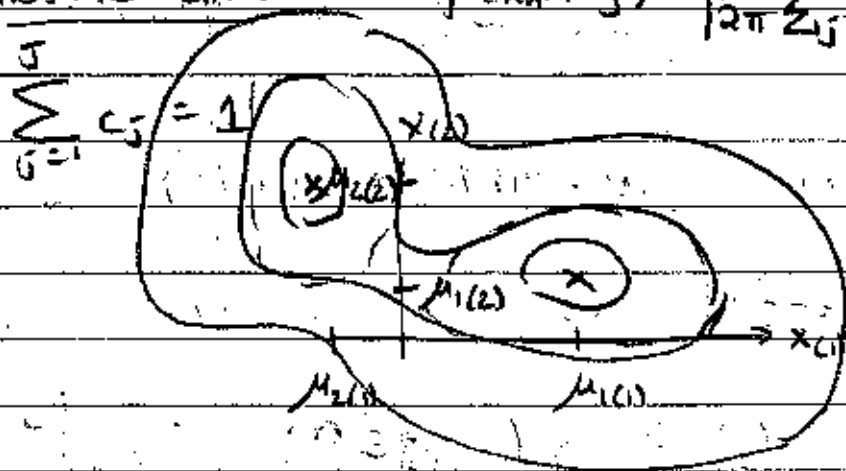


HOW TO REPRESENT



ANSWER:
$$p(\vec{x}_n | \theta) = \sum_{j=1}^J c_j p(\vec{x}_n | \theta_j)$$

MIXTURE GAUSSIAN:
$$p(\vec{x}_n | \theta_j) = \frac{1}{|2\pi \Sigma_j|^{1/2}} e^{-\frac{1}{2}(\vec{x}_n - \mu_j)^T \Sigma_j^{-1} (\vec{x}_n - \mu_j)}$$



ESTIMATION $G(\mathcal{X}, \theta) =$

$$\ln p(\mathcal{X} | \theta) = \sum_{n=1}^N \ln p(x_n | \theta)$$

$$= \sum_{n=1}^N \ln \left(\sum_{j=1}^J c_j p(x_n | \theta_{j,j}) \right)$$

$$\frac{\partial G(\mathcal{X}, \theta)}{\partial \theta} = 0$$

ALMOST GIVES A

SOLUTION, BUT NOT QUITE

II. VARIATIONAL BAYES $\sum_{j=1}^J q(j) = 1$

TRICK: INTRODUCE $q(j)$ = PSEUDO-DISTRIBUTION

$$\ln p(\mathcal{X} | \theta) = \sum_{n=1}^N \ln \left(\sum_{j=1}^J q(j) \frac{p(\vec{x}_n, j | \theta)}{q(j)} \right)$$

JENSEN'S INEQUALITY: SINCE $\sum_{j=1}^J q(j) = 1$,

$$\ln \left(\sum_{j=1}^J q(j) f_j \right) \geq \sum_{j=1}^J q(j) \ln f_j$$

SO VARIATIONAL BAYES

$$\ln p(\mathcal{X} | \theta) \geq \sum_{j=1}^J \sum_{n=1}^N q(j) \ln \left(\frac{c_j p(\vec{x}_n | \theta, j)}{q(j)} \right)$$

EXPECTATION MAXIMIZATION (EM)

① E-STEP FIND $q(j)$ TO MAXIMIZE RHS

$$\Rightarrow \textcircled{a} q(j) \propto p(\vec{x}_n, j | \theta) = P(\{j, \vec{x}_n | \theta\})$$

BUT

$$\textcircled{b} \sum_{j=1}^J q(j) = 1$$

$$\textcircled{a} + \textcircled{b} \Rightarrow q(j) = p(j | \vec{x}_n, \theta) = \frac{p(\vec{x}_n, j | \theta)}{\sum_{j=1}^J p(\vec{x}_n, j | \theta)}$$

② M-STEP FIND $\hat{\theta}$ TO MAXIMIZE RHS

$$\Rightarrow \hat{\theta} = \operatorname{argmax}_{\theta} \sum_{j=1}^J \sum_{n=1}^N q(j) \ln p(\vec{x}_n, j | \theta)$$

EQUIVALENTLY

① E-STEP COMPUTE $p(j | \vec{x}_n, \theta)$

② M-STEP FIND

$$\hat{\theta} = \operatorname{argmax}_{\theta} Q(\theta, \hat{\theta})$$

$$= \operatorname{argmax}_{\theta} \sum_{n=1}^N \left[E_j \left[\ln p(\vec{x}_n, j | \theta) \mid \theta \right] \right]$$

GAUSSIAN MIXTURE MODEL

① E-STEP $N(\vec{x}_n | \vec{\mu}_j, \Sigma_j) = \frac{1}{|2\pi\Sigma_j|^{1/2}} e^{-\frac{1}{2}(\vec{x}_n - \vec{\mu}_j)^T \Sigma_j^{-1} (\vec{x}_n - \vec{\mu}_j)}$

$$\gamma_n(j) = \frac{c_j N(\vec{x}_n | \vec{\mu}_j, \Sigma_j)}{\sum_{k=1}^J c_k N(\vec{x}_n | \vec{\mu}_k, \Sigma_k)} = p(j | \vec{x}_n, \theta)$$

② M-STEP

$$\hat{c}_j = \frac{1}{N} \sum_{n=1}^N \gamma_n(j)$$

$$\hat{\mu}_j = \frac{\sum_{n=1}^N \gamma_n(j) \vec{x}_n}{\sum_{n=1}^N \gamma_n(j)}$$

$$\hat{\Sigma}_j = \frac{\sum_{n=1}^N \gamma_n(j) (\vec{x}_n - \hat{\mu}_j)(\vec{x}_n - \hat{\mu}_j)^T}{\sum_{n=1}^N \gamma_n(j)}$$