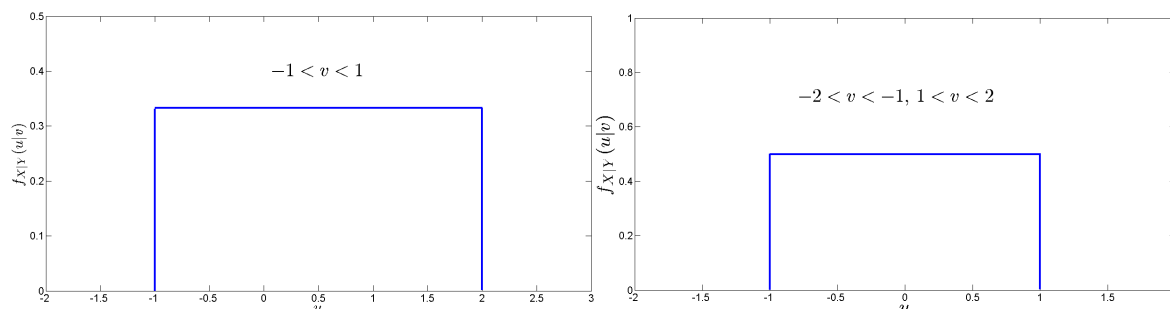


## ECE 313: Solutions to Problem Set 12

1. (a) For  $1 < |v| < 2$ , the cross-section of the joint pdf surface along the line  $v$  is a rectangle extending from  $u = -1$  to  $u = 1$ . Therefore the conditional pdf of  $X$  given  $Y$  is  $\text{Uniform}[-1, 1]$ .

Similarly for  $|v| < 1$ , the cross-section extends from  $u = -1$  to  $u = 2$ , so the conditional pdf of  $X$  given  $Y$  is  $\text{Uniform}[-1, 2]$ .



- (b) Using the uniformity of the conditional pdfs, we have:

$$\mathbb{E}[X|Y = v] = \begin{cases} 0, & 1 \leq |v| \leq 2 \\ \frac{1}{2}, & |v| \leq 1 \end{cases} = \text{rect}(v/2), \quad \text{Var}[X|Y = v] = \begin{cases} 1/3, & 1 \leq |v| \leq 2 \\ 3/4, & |v| \leq 1 \end{cases}$$

- (c) It is easy to see that  $f_Y(v) = \begin{cases} 0.2, & 1 < |v| < 2, \\ 0.3, & |v| < 1, \\ 0, & \text{otherwise.} \end{cases}$  Hence,

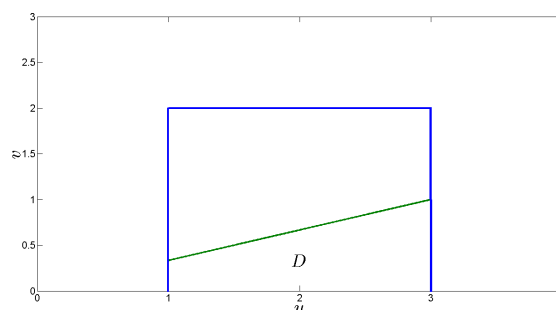
$$\mathbb{E}[X] = \int_{-\infty}^{\infty} \mathbb{E}[X|Y = v] \cdot f_Y(v) dv = \int_{-\infty}^{\infty} \text{rect}(v/2) \cdot f_Y(v) dv = \int_{-1}^1 \left(\frac{1}{2}\right) \cdot (0.3) dv = \frac{3}{10}$$

2. (a) First, note that the support of  $f_{X,Y}(u, v)$  is a square and the joint density does not depend on  $v$ . Thus, conditioned on any value of  $u$ , we see that  $f_{Y|X}(v|u)$  must be uniform on  $[0, 2]$ .

$$f_Y(v) = \int_{-\infty}^{\infty} f_{X,Y}(u, v) du = \int_1^2 \left(\frac{u}{8}\right) du = \frac{1}{2}, \quad 0 \leq v \leq 2 \Rightarrow f_Y(v) \sim \text{Uniform}[0, 2]$$

Since  $f_{Y|X}(v|u) = f_Y(v)$ ,  $X$  and  $Y$  are independent.

However, conditioned on  $D$ ,  $X$  and  $Y$  are not independent. Conditioned on  $D$ , the support of  $f_{X,Y|D}(u, v|D)$  is a trapezoid. We get this by intersecting the previous support (the rectangle) with the event  $\{v \leq \frac{u}{3}\}$  (see the diagram below). So, given  $D$ , the support of  $Y$  is  $[0, 1]$ . But given  $D$  and  $\{X = u\}$ , the support of  $Y$  is  $[0, \frac{u}{3}]$ , which depends on  $u$ . Thus, given  $D$ , the conditional density of  $Y$  given  $X$  and the marginal density of  $Y$  differ.

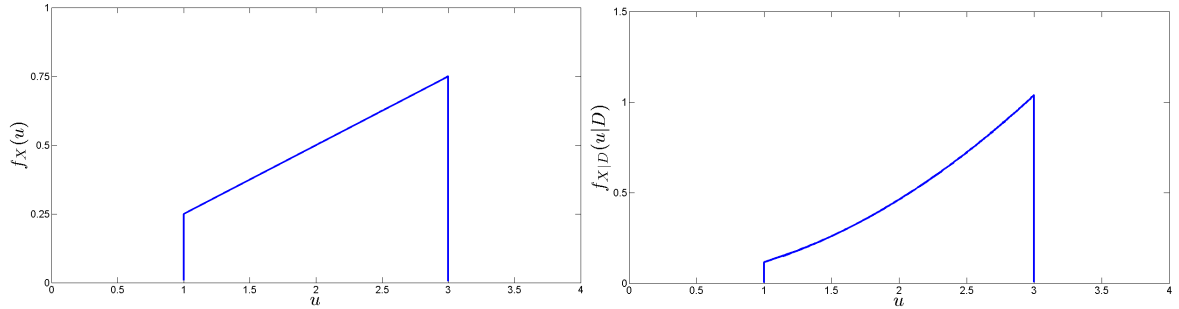


$$(b) \quad P(D) = P(R \leq 1/3) = P\left(\frac{Y}{X} \leq \frac{1}{3}\right) = P\left(Y \leq \frac{X}{3}\right) = \int_1^3 \int_{v=0}^{u/3} \left(\frac{u}{8}\right) dv du = \frac{13}{36}$$

$$(c) \quad \text{For } 1 \leq u \leq 3: f_X(u) = \int_{-\infty}^{\infty} f_{X,Y}(u,v) dv = \int_{v=0}^2 \left(\frac{u}{8}\right) dv = \frac{u}{4}$$

$$\text{For } 1 \leq u \leq 3: \int_{(u,v) \in D} f_{X,Y}(u,v) dv = \int_0^{u/3} \left(\frac{u}{8}\right) dv = \frac{u^2}{24}$$

$$f_{X|D}(u|D) = \frac{\int_{(u,v) \in D} f_{X,Y}(u,v) dv}{P(D)} = \frac{u^2/24}{13/36} = \begin{cases} \frac{3u^2}{26}, & 1 \leq u \leq 3 \\ 0, & \text{otherwise} \end{cases}$$



$$(d) \quad F_{R|D}(r|D) = P\{R \leq r|D\} = \frac{P\{R \leq r, D\}}{P\{D\}}$$

$$0 \leq r \leq \frac{1}{3} : P\{R \leq r \cap D\} = P\{R \leq r\} = P\{Y \leq rX\} = \int_1^3 \int_0^{ru} \left(\frac{u}{8}\right) dv du = \frac{13r}{12}$$

$$r \geq \frac{1}{3} : P\{R \leq r \cap D\} = P\left\{R \leq \frac{1}{3}\right\} = P\{D\} = \frac{13}{36}$$

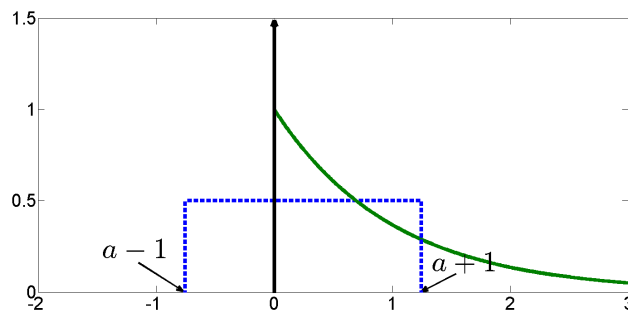
$$F_{R|D}(r|D) = \begin{cases} 0, & r \leq 0 \\ 3r, & 0 \leq r \leq \frac{1}{3} \\ 1, & r \geq \frac{1}{3} \end{cases} \Rightarrow f_{R|D}(r|D) = \begin{cases} 3, & 0 \leq r \leq \frac{1}{3} \\ 0, & \text{otherwise} \end{cases}$$

Thus, succinctly, conditioned on  $D$ ,  $R$  is uniformly distributed on  $[0, 1/3]$ .

3. The densities of  $X_1$  and  $X_2$  are given by:

$$f_{X_1}(u) = \begin{cases} \frac{1}{2} & -1 \leq u \leq 1 \\ 0 & \text{elsewhere} \end{cases}, \quad f_{X_2}(u) = \begin{cases} e^{-u} & u \geq 0 \\ 0 & u < 0 \end{cases}$$

pdf of  $W = X_1 + X_2$ : Since  $X_1$  and  $X_2$  are independent, the pdf of  $W$  is just the convolution of the individual pdfs (here, we are keeping the exponential density fixed and flipping and shifting the uniform density – see diagram below).

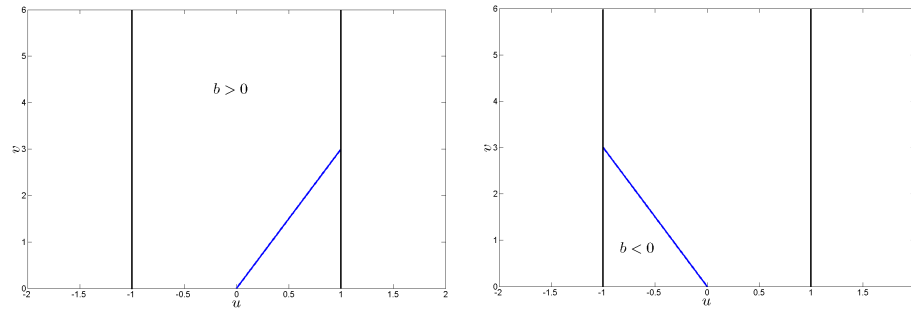


$$f_W(a) = \begin{cases} 0, & a < -1 \\ \int_0^{a+1} \frac{1}{2} e^{-u} du, & -1 \leq a < 1 \\ \int_{a-1}^{a+1} \frac{1}{2} e^{-u} du, & a > 1 \end{cases} = \begin{cases} 0, & a < -1 \\ \frac{1}{2} (1 - e^{-(a+1)}), & -1 \leq a < 1 \\ \frac{1}{2} (e - e^{-1}) e^{-a}, & a > 1 \end{cases}$$

pdf of  $Z = X_1/X_2$ : Since  $X_2 \geq 0$ ,

$$\begin{aligned} F_Z(b) = P\{X_1/X_2 \leq b\} &= P\{X_1 \leq bX_2\} = \begin{cases} \int_{u=-1}^0 \int_{v=0}^{u/b} \frac{1}{2} e^{-v} dv du, & b < 0 \\ 1 - \int_{u=0}^1 \int_{v=0}^{u/b} \frac{1}{2} e^{-v} dv du, & b \geq 0 \end{cases} \\ &= \begin{cases} \frac{1}{2} (1 + b(1 - e^{1/b})), & b < 0 \\ \frac{1}{2} (1 + b(1 - e^{-1/b})), & b \geq 0 \end{cases} \end{aligned}$$

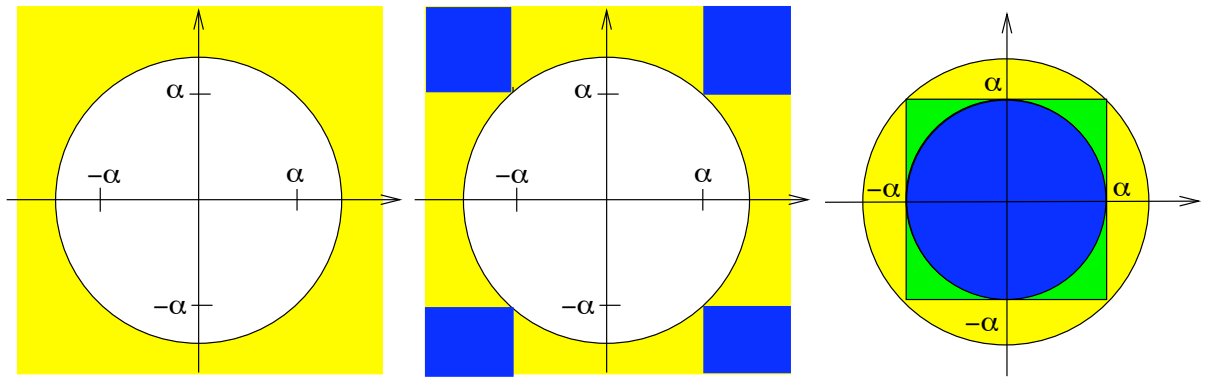
The areas over which to integrate are indicated in the figures below.



From this, the pdf of  $Z$  can be obtained by differentiating

$$f_Z(b) = \begin{cases} \frac{1}{2} \left( 1 - e^{1/b} + \frac{e^{1/b}}{b} \right), & b < 0 \\ \frac{1}{2} \left( 1 - e^{-1/b} - \frac{e^{-1/b}}{b} \right), & b \geq 0 \end{cases}$$

4. (a)  $f_{\mathcal{X}, \mathcal{Y}}(u, v) = f_{\mathcal{X}}(u) f_{\mathcal{Y}}(v) = \frac{1}{2\pi} \exp \left[ -\frac{u^2 + v^2}{2} \right]$ .
- (b) The region over which the joint pdf must be integrated in order to find  $P\{\mathcal{X}^2 + \mathcal{Y}^2 > 2\alpha^2\}$  is shown in the left-hand figure below.



The volume outside the circle of radius  $\sqrt{2}\alpha$  can be found by changing to polar coordinates. We have  $P\{\mathcal{X}^2 + \mathcal{Y}^2 > 2\alpha^2\} = \int_{r=\sqrt{2}\alpha}^{\infty} \int_{\theta=0}^{2\pi} \frac{1}{2\pi} \exp \left( -\frac{r^2}{2} \right) r d\theta dr = \int_{r=\sqrt{2}\alpha}^{\infty} r \exp \left( -\frac{r^2}{2} \right) dr$   
 $= -\exp \left( -\frac{r^2}{2} \right) \Big|_{\sqrt{2}\alpha}^{\infty} = \exp(-\alpha^2).$

- (c)  $\mathcal{Z} = \mathcal{X}^2 + \mathcal{Y}^2$ . From part (b), we get that  $P\{\mathcal{Z} > \beta\} = 1 - F_{\mathcal{Z}}(\beta) = \exp(-\beta/2)$  for  $\beta \geq 0$ . Hence,  $f_{\mathcal{Z}}(\beta) = \frac{1}{2} \exp(-\beta/2)$  for  $\beta \geq 0$ . This is an exponential density with parameter  $\frac{1}{2}$ .
- (d)  $P\{|\mathcal{X}| > \alpha\} = 2Q(\alpha)$ . Hence,  $P\{|\mathcal{X}| > \alpha, |\mathcal{Y}| > \alpha\} = P\{|\mathcal{X}| > \alpha\}P\{|\mathcal{Y}| > \alpha\} = 4Q^2(\alpha)$ .
- (e) From the middle figure above, it is obvious that

$$P\{|\mathcal{X}| > \alpha, |\mathcal{Y}| > \alpha\} = 4Q^2(\alpha) < P\{\mathcal{X}^2 + \mathcal{Y}^2 > 2\alpha^2\} = \exp(-\alpha^2) \text{ for } \alpha > 0.$$

- (f) Taking square roots on both sides, we get  $Q(x) < \frac{1}{2} \exp(-x^2/2)$  for  $x > 0$ .
- (g)  $P\{|\mathcal{X}| < \alpha, |\mathcal{Y}| < \alpha\} = P\{|\mathcal{X}| < \alpha\}P\{|\mathcal{Y}| < \alpha\} = [1 - 2Q(\alpha)]^2$  is the probability that the random point lies inside the square of side  $2\alpha$ . As shown in the right-hand figure, this square is inscribed by the circle of radius  $\alpha$  and circumscribed by the circle of radius  $\sqrt{2}\alpha$ . Hence,

$$P\{\mathcal{X}^2 + \mathcal{Y}^2 \leq \alpha^2/2\} < P\{|\mathcal{X}| < \alpha, |\mathcal{Y}| < \alpha\} < P\{\mathcal{X}^2 + \mathcal{Y}^2 < \alpha^2\}.$$

These inequalities are equivalent to the result:  $\exp(-\alpha^2) < 4Q(\alpha) - 4Q^2(\alpha) < \exp(-\alpha^2/2)$ . Now,  $4Q(\alpha) - 4Q^2(\alpha) < 4Q(\alpha)$ , and therefore we get  $Q(x) > \frac{1}{4} \exp(-x^2)$  for  $x > 0$ . Note that  $Q(0) = \frac{1}{2}$  while the bound equals  $\frac{1}{4}$  at  $x = 0$ . The upper bound can also be obtained from the above result. Note that since  $Q(\alpha) \leq \frac{1}{2}$  for  $\alpha > 0$ , it follows that  $4Q(\alpha) - 4Q^2(\alpha) = 4Q(\alpha)[1 - Q(\alpha)] > 2Q(\alpha)$  for  $\alpha > 0$ . Hence,  $\exp(-\alpha^2/2) > 4Q(\alpha) - 4Q^2(\alpha) > 2Q(\alpha)$  etc.

5. (a) Converting to polar coordinates:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(u,v) du dv = \int_{\theta=0}^{2\pi} \int_{r=0}^1 cr \cdot r dr d\theta = \frac{2\pi c}{3} = 1 \Rightarrow c = \frac{3}{2\pi}$$

- (b) Using LOTUS:

$$\mathbb{E}[f(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{3}{2\pi} \sqrt{u^2 + v^2} \cdot \frac{3}{2\pi} \sqrt{u^2 + v^2} du dv = \frac{9}{4\pi^2} \int_{\theta=0}^{2\pi} \int_{r=0}^1 r^3 dr d\theta = \frac{9}{8\pi}$$

- (c)  $f(X,Y) = \frac{3}{2\pi} \sqrt{X^2 + Y^2}$  has maximum value  $\frac{3}{2\pi} < \frac{1}{2}$ . Therefore  $P\{f(X,Y) > \frac{1}{2}\} = 0$ .