

ECE 313
Probability with Engineering Applications
Lecture 38 – Conditional Distributions
Professor Seth Hutchinson
for Dilip V. Sarwate
 Department of Electrical and
 Computer Engineering

© 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved

Last Time

Last Time (expectation)
 $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X,Y)$
 $\rho(X,Y) = \frac{\text{Cov}(X,Y)}{\sigma_X \sigma_Y}, -1 \leq \rho \leq 1$

$\rho = 0 \rightarrow X+Y$ uncorrelated

Independent $\begin{matrix} \rightarrow \\ \leftarrow \end{matrix}$ Uncorrelated

ECE 313 - Lecture 38 © 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved Slide 1 of 3

Jointly Gaussian rv's

Jointly Gaussian rv's

- ρ appears in $f_{X,Y}$, but f_X, f_Y do not depend on ρ
- $\rho = 0 \Rightarrow$ Independence !!!!!
- $X+Y$ Jointly Gauss. $\begin{matrix} \rightarrow \\ \leftarrow \end{matrix}$ X Gauss. Y Gauss.
- Level sets of $f_{X,Y}$ are ellipses.

ECE 313 - Lecture 38 © 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved Slide 2 of 3

Sum of Jointly Gaussian rv's

Sum of Jointly Gaussian rv's

- if $X+Y$ J.G. + Independent $Z = X+Y$
 $Z \sim \mathcal{N}(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2)$
- If $X+Y$ not independent...
 • rotate axes $\Rightarrow \rho = 0 \Rightarrow W+Z$ ind.
 • Sum must be Gaussian r.v.

$E[Z] = E[X+Y] = E[X] + E[Y]$
 $\Rightarrow \mu_Z = \mu_X + \mu_Y$

$\text{Var}(Z) = \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X,Y)$
 Exercise: $W = aX + bY \Rightarrow$ find μ_W, σ_W^2

ECE 313 - Lecture 38 © 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved Slide 3 of 3

Conditional Distributions

Conditional Distributions To Date:

- For events: $P(B) > 0$
 $P(A|B) = \frac{P(A \cap B)}{P(B)}$
- For discrete rv's: $P_X(a) > 0: P_{X|Y}(u|v) = \frac{P_{X,Y}(u,v)}{P_Y(v)}$
- For continuous rv X : $P_X(a) > 0$
 $f_{X|A}(u|A) = P\{X \leq u | A\} = \frac{P\{X \leq u, A\}}{P\{A\}}$
 $f_{X|A} = \frac{d}{du} F_{X|A}$

Suppose we observe $\{Y=v\}$ - how to condition on this

ECE 313 - Lecture 38 © 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved Slide 4 of 3

Recall...

Recall: The probability that (X,Y) lies in a differential region is:

$P\{u \leq X \leq u+du, v \leq Y \leq v+dv\} \approx f_{X,Y}(u,v) du dv$

And: $P\{v \leq Y \leq v+dv\} \approx f_Y(v) dv \neq 0$

Thus:
 $P\{u \leq X \leq u+du | v \leq Y \leq v+dv\} = \frac{P\{u \leq X \leq u+du, v \leq Y \leq v+dv\}}{P\{v \leq Y \leq v+dv\}}$

Doesn't depend on $dv \Rightarrow \approx \frac{f_{X,Y}(u,v) du dv}{f_Y(v) dv}$
 $dv \rightarrow 0$

$P\{u \leq X \leq u+du | Y=v\} \approx \frac{f_{X,Y}(u,v) du}{f_Y(v)}$

ECE 313 - Lecture 38 © 2000 Seth Hutchinson, University of Illinois at Urbana-Champaign. All Rights Reserved Slide 5 of 3

$$f_{X|Y}(u|v) du = \frac{f_{XY}(u,v) du}{f_Y(v)}$$

⇒ Define $f_{X|Y}(u|v) = \frac{f_{XY}(u,v)}{f_Y(v)}$

for $f_Y(v) \neq 0$

Notation $f_{X|Y}(u|v)$

↑ known
Variable

$f_{X|Y}(u|3/4)$

ECE 313 - Lecture 38 © 2000 Scott Holman, University of Illinois at Urbana-Champaign, All Rights Reserved Slide 6 of 8

Example

EX. $f_{XY}(u,v) = \begin{cases} 2: & 0 < u < v < 1 \\ 0: & \text{else} \end{cases}$

→ $f_X(u) = 2(1-u) \quad 0 < u < 1$

→ $f_{Y|X}(v|u) = \frac{2}{2(1-u)} = \frac{1}{1-u} \quad u < v < 1$

uniform on $(u, 1)$

ECE 313 - Lecture 38 © 2000 Scott Holman, University of Illinois at Urbana-Champaign, All Rights Reserved Slide 7 of 8

Example

EX. $X+Y$ Joint. Gauss. $\mu_X = \mu_Y = 0, \sigma_X^2 = \sigma_Y^2 = 1$

$$f_{Y|X}(u|v) = \frac{f_{XY}(u,v)}{f_X(v)} = \frac{\frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} [u^2 - 2\rho uv + v^2]}}{\frac{1}{\sqrt{2\pi}} e^{-u^2/2}}$$

$$= \frac{1}{\sqrt{2\pi}\sqrt{1-\rho^2}} e^{-\frac{(v-\rho u)^2}{2(1-\rho^2)}}$$

$\mathcal{N}(\rho u, 1-\rho^2)$

ECE 313 - Lecture 38 © 2000 Scott Holman, University of Illinois at Urbana-Champaign, All Rights Reserved Slide 8 of 8