

Review of Lecture #3

- We explored the consequences of the axioms of probability
- We discussed the notion of a partition and how partitions can be used to calculate the probability of an event
- We discussed binomial coefficients
- We looked at a simple combinatorial problem

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Drawing a random sample

- Example: An urn contains 6 identical red balls $R_1, R_2, R_3, R_4, R_5, R_6$ and 3 identical green balls G_1, G_2, G_3 . A trial of the experiment consists of **simultaneously** drawing **two balls at random** from the urn
- The **outcomes** of this experiment are **subsets of size 2** of the form $\{R_1, R_5\}$ or $\{R_4, G_1\}$ or $\{G_2, G_3\}$

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How many subsets of each kind?

- There are $\binom{9}{2} = 36$ subsets of size 2 from a set of 9 balls, and the collection of these subsets is the sample space
- $\binom{6}{2} = 15$ outcomes consist of two red balls
- $\binom{3}{2} = 3$ outcomes have two green balls
- $\binom{6}{1} \binom{3}{1} = 18$ outcomes have 1 red, 1 green

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More probabilities

- $A = \{R_2 \text{ is in the sample drawn}\}$
- $B = \{G_2 \text{ is in the sample drawn}\}$
- What is $P(A \cap B)$?
- $P(A) = 8/36$ $P(B) = 8/36$
- $AB = \{\text{outcome} = \{R_2, G_2\}\}$ is a singleton event, and hence $P(AB) = 1/36$
- $P(A \cup B) = P(A) + P(B) - P(AB) = 15/36$
- Exercise: What is $P(A^c \cap B^c)$?

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Sampling without replacement

- A sample of size k from a set of size n is a subset of size k
- $\{R_2, G_2\}$ is the same as $\{G_2, R_2\}$
- Instead of obtaining the k elements of the subset simultaneously, we could draw them out one at a time, **each draw being carried out without replacing the previously drawn elements**
- This is **sampling without replacement**

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Sampling without replacement

- If a previously drawn element were to be put back, then it **could** be drawn again
- This is not what we want!
- Sampling without replacement results in the same subsets of size k as random sampling, but the elements in the subset are arranged in a specific order
- Outcomes should be viewed as **vectors**: (R_2, G_2) is not the same as (G_2, R_2)

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Sampling without replacement

- A subset of size k can be arranged into $k!$ different vectors
- Outcomes of the experiment are $k! \binom{n}{k}$
 $= n(n-1)(n-2)\dots(n-k+1)$ vectors of length k whose entries are k distinct elements from the set of size n
- Usual explanation: n choices for first entry, then $n-1$ for second, $n-2$ for third, etc.

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Sampling without replacement

- Example: Our urn contains 6 identical red balls $R_1, R_2, R_3, R_4, R_5, R_6$ and 3 identical green balls G_1, G_2, G_3 . Two balls are drawn at random without replacement from the urn
- 72 different samples without replacement can be drawn from the urn
- As usual, at random means that each of the 72 outcomes is equally likely

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A surprising result?

- $P\{\text{first ball is green}\}$
 $= P\{\text{outcome is } (G_1, *)\}$
 $+ P\{\text{outcome is } (G_2, *)\}$
 $+ P\{\text{outcome is } (G_3, *)\}$
 $= (8 + 8 + 8)/72 = 1/3$
- $P\{\text{second ball drawn is green}\} = P\{(*, G_1)\}$
 $+ P\{(*, G_2)\} + P\{(*, G_3)\} = 1/3$ also!
- Here, $*$ and $*$ denote any ball **other** than the G_i that is exhibited

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Surprising result is quite general

- More specifically, $P\{\text{first ball drawn is } G_2\}$
 $= P\{\text{outcome is } (G_2, *)\} = 8/72 = 1/9$
- $P\{\text{second ball drawn is } G_2\} = P\{(*, G_2)\}$
 $8/72 = 1/9$ also!
- More generally, if (x_1, x_2, \dots, x_k) is the random sample without replacement, and A is some subset (e.g. "green ball" or $\{G_2\}$) of the set of n elements, then
 $P\{x_1 \in A\} = P\{x_2 \in A\} = \dots = P\{x_k \in A\}$

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$P\{\text{both balls are green}\}$?

- $P\{\text{both balls are green}\}$
 $= P\{(G_1, G_2), (G_2, G_1), (G_1, G_3), (G_3, G_1),$
 $(G_2, G_3), (G_3, G_2)\}$
 $= 6/72 = 3/36$ as with a random sample
- Sampling without replacement is a "closer look" at a random sample: the ordering is important in the former, not in the latter
- When the ordering is irrelevant, as in {both balls are green}, we should get the same result, and we did

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Problems with classical approach

- The classical approach of equally likely outcomes cannot be applied to sample spaces with infinitely many outcomes
- The nonclassical approach **does** extend to some, but not all types, of infinite sample spaces
- Can a sample space have infinitely many outcomes?
- **Yes**

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Infinitely many outcomes

- Example: The experiment consists of tossing a coin till a Tail appears for the first time. The sample space is $= \{T, HT, HHT, HHHT, \dots\}$
- This sample space is **countably infinite**
- **Countable** means there is a **one-to-one correspondence** between the integers and the outcomes

Integer n $n-1$ heads
 HHH...HT

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It's not in my job description!

- **Automatic Repeat Request (ARQ)** communication systems transmit a data packet (with error-detection mechanism such as a CRC checksum)
- The receiver replies with a NACK or ACK according as it detects errors or not
- Sender **repeats** transmission if the receiver replies NACK
- Do it over and over again till it's done right!

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Countably infinite sample spaces

- Let $\Omega = \{x_1, x_2, \dots, x_n, \dots\}$ be the countably infinite sample space
- $P\{x_n\} = p_n$ where $p_n \geq 0$
- For a finite subset A of Ω , $P(A)$ is just the sum of the probabilities of the outcomes comprising the event A , as before
- It seems reasonable to have this idea work for an infinite subset of A as well
- But we need a new improved axiom

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Axiom III needs improvement

- Axiom III for a finite sample space states that $P(A \cup B) = P(A) + P(B)$ for disjoint events A and B
- For a **finite** set of disjoint events, this readily extends to $P(A \cup B \cup \dots \cup G) = P(A) + P(B) + \dots + P(G)$, but it **does not extend** (e.g. via mathematical induction) to an **infinite collection** of disjoint events
- Mathematical details are too technical

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New improved Axiom III

- **New improved Axiom III:** Let $A_1, A_2, \dots, A_n, \dots$ denote a **countable sequence** of **disjoint** events, that is, $A_i \cap A_j = \emptyset$ for all $i \neq j$. Then, $P(A_1 \cup A_2 \cup \dots \cup A_n \cup \dots) = P(A_1) + P(A_2) + \dots + P(A_n) + \dots$
- The new Axiom III implies that $P(\emptyset) = 0$ and $P(A \cup B) = P(A) + P(B)$ for $A \cap B = \emptyset$
- See Ross, p. 31 for details

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Better living through new Axiom III

- For a finite subset A of Ω , $P(A)$ is just the sum of the probabilities of the outcomes comprising the event A , as before
- The new improved Axiom III extends this to infinite subsets as well
- With the new improved Axiom III, we can treat countably infinite sample spaces just as we handle finite sample spaces

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Some thoughts about the p_n

- Special case: $\{x_1, x_2, \dots, x_n, \dots\}$
 $= \{x_1\} \{x_2\} \dots \{x_n\} \dots$ and hence
 $P(\cdot) = P\{x_1\} + P\{x_2\} + \dots + P\{x_n\} + \dots$
 $= p_1 + p_2 + \dots + p_n + \dots = 1$
- The p_i 's cannot all be equal (Why not?)
- The sequence of p_n 's converges to 0
- The series $p_1 + p_2 + \dots + p_n + \dots$ converges to 1, and any **subseries** of this series also converges

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An example

- $\{x_1, x_2, \dots, x_n, \dots\}$ where
 $p_n = P\{x_n\} = (0.6) \times (0.4)^{n-1}$ for all n
- $p_1 + p_2 + \dots + p_n + \dots$
 $= (0.6) \times [1 + (0.4) + \dots + (0.4)^{n-1} + \dots]$
 $= (0.6) \times 1/[1 - 0.4] = 1$ on summing the geometric series
- $1/[1-x] = 1 + x + x^2 + \dots + x^n + \dots$
 for $|x| < 1$ is one of the most useful results in this course. **Memorize it!**

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Example (continued)

- $\{x_1, x_2, \dots, x_n, \dots\}$ where
 $p_n = P\{x_n\} = (0.6) \times (0.4)^{n-1}$ for all n
- Let $A = \{x_2, x_4, \dots, x_{2k}, \dots\}$ denote the event that the outcome of the experiment is x_i where i is an **even** integer
- $P(A) = p_2 + p_4 + \dots + p_{2k} + \dots$
 $= (0.6) \times (0.4) [1 + (0.4)^2 + (0.4)^4 + \dots]$
 $= (0.24) \times 1/[1 - (0.4)^2] = 2/7$

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What do probabilities mean?

- The classical approach to probability justifies its equally likely outcomes with an appeal to symmetry
- The nonclassical approach does not necessarily provide reasons for its assignments of probabilities
- So, **why** should different outcomes have different probabilities? What is the reason for setting $P(x) = 0.1$ and $P(y) = 0.2$?

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Observed relative frequencies

- Probabilities are expressions of belief about what is going to happen in the future
- One reason for setting $P(x) = 0.1$ and $P(y) = 0.2$ might be based on **what was observed in the past**
- If outcome x was observed N_x times on N trials of the experiment, then N_x/N is called the **observed relative frequency** of x
- For events: $N_A = \#$ of occurrences of A

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Relative frequency = probability?

- **Observed relative frequencies** satisfy the probability axioms
- Event A was observed N_A times on N trials where $N_A \geq 0$. Hence, $N_A/N \geq 0$
- x was observed on all trials: $N_x/N = 1$
- If A and B are disjoint events, then on the trials on which A occurred, B could not have occurred, and vice versa. Hence,
 $N_{A \cup B} = N_A + N_B$ $N_{A \cup B}/N = N_A/N + N_B/N$

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Estimates of probabilities

- Based on past experimental results, we can use the observed relative frequency as an **estimate** of the probability of an event
- Estimates are **usually** reasonably good if the number of trials was large
- On N trials, the relative frequency that we observe **might** easily differ from the **true probability** by as much as $N^{-1/2}$ or more

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But they don't always work!

- For a sample space with countably infinite outcomes, at most N different outcomes will be observed on N trials
- The probabilities to be assigned to the infinitely many **unobserved** outcomes is again a matter of guesswork!
- Since the sequence of p_n 's converges to 0, most of the probabilities are much smaller than $N^{-1/2}$. Estimation errors?

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A million tosses, a million laughs

- Suppose that a **fair** coin is tossed a million times
- Is there a **logical** reason why the coin will not turn up Heads **each and every** time?
- **No**, there is no logical reason why it couldn't, **but it is very unlikely to do so**
- **Yes**, if the coin is fair, there is **no way** that it can turn up Heads a million times in a row

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Faith versus belief

- A **fair** coin that is tossed a million times will not turn up Heads a million times in a row?
- What is the largest number of consecutive Heads that you think can possibly occur?
- How does the coin know that you are watching, and therefore, it should turn up Tails after a certain number of Heads?
- What if more than one person is watching?

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No guarantees

- A million Heads in a row might later be followed by a million Tails to give a relative frequency of $1/2$ for Heads
- If only we had gone one for another million tosses, we would have been OK ...
- No **guarantee** that an observed relative frequency is close to the actual probability
- We **expect** the relative frequency is close, but we cannot guarantee it

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Estimates only

- Observed relative frequencies can serve only as **estimates** of probabilities
- Do not use them to **define** probabilities even though they satisfy the axioms
- No **guarantee** that an observed relative frequency is close to the actual probability, only a strong **expectation**
- "What does probability mean?" It is a numerical expression of strength of belief

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Large number of trials

- In a **large number** of trials (say coin tosses) **long** runs of Heads (or Tails) can be expected to occur
- Actual coin tosses vs human simulation
- The **gambler's ruin**: If #Heads > #Tails (say) at any particular point in the game, then this lead will last for a **long time**
- **Lead changes** become **more infrequent**
- Deviations increase in absolute value

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Limit of relative frequency?

- A lot of effort was expended in trying to **define** probability as the limit of the relative frequency
- $P(A) = \lim_N N_A/N$
- Unfortunately, the limit does not exist in a mathematical sense
- Physically we will only observe a finite-length prologue of the sequence of trials

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Uncountably infinite sample spaces

- The nonclassical approach cannot be used at all for **uncountably infinite** sample spaces
- Example: Pick a **random** number between 0 and 1
- $= \{x : 0 < x < 1\}$
- Intuitively, the meaning of random in this instance is that we do not favor any one number in the interval (0,1) over the others

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Example (continued)

- Example: Pick a **random** number between 0 and 1
- $= \{x : 0 < x < 1\} = (0, 1)$
- One way of expressing the innate randomness of the choice is as follows:
- Given any subinterval of (0, 1), the probability that the chosen number lies in that subinterval is equal to the length of the interval

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God made the integers

- Kronecker: God made the integers; all else is the work of man
- Human beings usually choose rational numbers when asked for a number in (0,1)
- A physical measurement made with an instrument will yield a rational number
- `rand()` returns "real numbers" that are actually rational numbers
- All this is because of finite precision

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Do real numbers exist?

- The real number line is a mathematical construction that models the real world very well indeed
- If the volume electrical charge density is ρ , the charge in a volume v is just ρv
- For v very small, ρv is smaller than the charge of an electron, so the model cannot be right for small volumes (or densities)!
- But it is convenient!

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Physical Science and Engineering

- In the physical sciences and engineering, the real numbers are a model for very many phenomena that are discrete at the microscopic level
- This usually causes no problems and the model usually gives the correct answers
- Calculus can be applied
- We all understand that $V = 1.235$ volts really means $1.2345 \leq V \leq 1.2355$ volts

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Real numbers in probability

- The uncountably infinite sample spaces that are intervals of the real line (or the entire real line) cause many **subtle mathematical difficulties** in probability theory
- There are some obvious problems as well
- The resolution of these difficulties led to the development of the axiomatic theory of probability

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What is $P\{\text{outcome}\}$?

- The output of `rand()` is a good model for repeated trials for the experiment of picking a number at random in $(0, 1)$
- Consider a million random numbers obtained from `rand()`. A particular outcome, say 0.703546789, will either not have occurred in these million trials, or it will have occurred just once.
- $P\{0.703546789\} = 10^{-6}$ or 0

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$P\{\text{outcome}\}$ is always 0?

- The relative frequency estimate of $P\{0.703546789\} = 10^{-6}$ or 0
- If `rand()` is invoked a further million times, then 0.703546789 will have occurred a total of 0 or 1 times in 2 million trials $P\{0.703546789\} = 0.5 \times 10^{-6}$ or 0
- The relative frequency estimate seems to be **converging** to 0 as the number of trials increases!

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$P\{\text{outcome}\}$ is always 0

- The **only** model that works for uncountably infinite sample spaces is for **each outcome to have probability 0**
- But, on each trial, **some** outcome occurs, doesn't it?
- So where are the probabilities?
- For `rand()`, $P\{a < \text{outcome} < b\} = b - a$
- The **nonzero probabilities** are assigned to the **intervals** of the line, **not to outcomes!**

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Summary

- We studied sampling without replacement and its relation to random samples
- We studied sample spaces with countably infinite outcomes and adopted a new Axiom III
- We waxed philosophical about relative frequencies, estimates, faiths and beliefs
- We began to study uncountably infinite sample spaces and noted $P\{\text{outcome}\} = 0$

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