## Chapter 23

## Entropy, Randomness, and Information

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### 23.1 Entropy

### 23.1.0.1 Quote

"If only once - only once - no matter where, no matter before what audience - I could better the record of the great Rastelli and juggle with thirteen balls, instead of my usual twelve, I would feel that I had truly accomplished something for my country. But I am not getting any younger, and although I am still at the peak of my powers there are moments - why deny it? - when I begin to doubt - and there is a time limit on all of us."
-Romain Gary, The talent scout.

### 23.2 Entropy

### 23.2.0.2 Entropy: Definition

Definition 23.2.1. The entropy in bits of a discrete random variable $X$ is

$$
\mathbb{H}(X)=-\sum_{x} \operatorname{Pr}[X=x] \lg \operatorname{Pr}[X=x]
$$

Equivalently, $\mathbb{H}(X)=\mathbf{E}\left[\lg \frac{1}{\operatorname{Pr}[X]}\right]$.

### 23.2.0.3 Entropy intuition...

Intuition... $\mathbb{H}(X)$ is the number of fair coin flips that one gets when getting the value of $X$.
Interpretation from last lecture... Consider a (huge) string $S=s_{1} s_{2} \ldots s_{n}$ formed by picking characters independently according to $X$. Then

$$
|S| \mathbb{H}(X)=n \mathbb{H}(X)
$$

is the minimum number of bits one needs to store the string $S$.

### 23.2.0.4 Binary entropy

$\xrightarrow{\mathbb{H}(X)=}-\sum_{x} \operatorname{Pr}[X=x] \lg \operatorname{Pr}[X=x]$
Definition 23.2.2. The binary entropy function $\mathbb{H}(p)$ for a random binary variable that is 1 with probability $p$, is $\mathbb{H}(p)=-p \lg p-(1-p) \lg (1-p)$. We define $\mathbb{H}(0)=\mathbb{H}(1)=0$.

Q: How many truly random bits are there when given the result of flipping a single coin with probability $p$ for heads?
23.2.0.5 Binary entropy: $\mathbb{H}(p)=-p \lg p-(1-p) \lg (1-p)$

(A) $\mathbb{H}(p)$ is a concave symmetric around $1 / 2$ on the interval $[0,1]$.
(B) maximum at $1 / 2$.
(C) $\mathbb{H}(3 / 4) \approx 0.8113$ and $\mathbb{H}(7 / 8) \approx 0.5436$.
(D) $\Longrightarrow$ coin that has $3 / 4$ probably to be heads have higher amount of "randomness" in it than a coin that has probability $7 / 8$ for heads.

### 23.2.0.6 And now for some unnecessary math

(A) $\mathbb{H}(p)=-p \lg p-(1-p) \lg (1-p)$
(B) $\mathbb{H}^{\prime}(p)=-\lg p+\lg (1-p)=\lg \frac{1-p}{p}$
(C) $\mathbb{H}^{\prime \prime}(p)=\frac{p}{1-p} \cdot\left(-\frac{1}{p^{2}}\right)=-\frac{1}{p(1-p)}$.
(D) $\Longrightarrow \mathbb{H}^{\prime \prime}(p) \leq 0$, for all $p \in(0,1)$, and the $\mathbb{H}(\cdot)$ is concave.
(E) $\mathbb{H}^{\prime}(1 / 2)=0 \Longrightarrow \mathbb{H}(1 / 2)=1$ max of binary entropy.
$(\mathrm{F}) \Longrightarrow$ balanced coin has the largest amount of randomness in it.

### 23.2.1 Task at hand: Squeezing good random bits...

### 23.2.1.1 ...out of bad random bits...

(A) $b_{1}, \ldots, b_{n}$ : result of $n$ coin flips...
(B) From a faulty coin!
(C) $p$ : probability for head.
(D) We need fair bit coins!
(E) Convert $b_{1}, \ldots, b_{n} \Longrightarrow b_{1}^{\prime}, \ldots, b_{m}^{\prime}$.
(F) New bits must be truly random: Probability for head is $1 / 2$.
(G) Q: How many truly random bits can we extract?

### 23.2.2 Intuitively...

### 23.2.2.1 Squeezing good random bits out of bad random bits...

Question... Given the result of $n$ coin flips: $b_{1}, \ldots, b_{n}$ from a faulty coin, with head with probability $p$, how many truly random bits can we extract?

If believe intuition about entropy, then this number should be $\approx n \mathbb{H}(p)$.

### 23.2.2.2 Back to Entropy

(A) entropy of $X$ is $\mathbb{H}(X)=-\sum_{x} \operatorname{Pr}[X=x] \lg \operatorname{Pr}[X=x]$.
(B) Entropy of uniform variable..

Example 23.2.3. A random variable $X$ that has probability $1 / n$ to be $i$, for $i=1, \ldots, n$, has entropy $\mathbb{H}(X)=-\sum_{i=1}^{n} \frac{1}{n} \lg \frac{1}{n}=\lg n$.
(C) Entropy is oblivious to the exact values random variable can have.
(D) $\Longrightarrow$ random variables over $-1,+1$ with equal probability has the same entropy (i.e., 1 ) as a fair coin.

### 23.2.2.3 Lemma: Entropy additive for independent variables

### 23.2.2.4 Lemma: Entropy additive for independent variables

Lemma 23.2.4. Let $X$ and $Y$ be two independent random variables, and let $Z$ be the random variable $(X, Y)$. Then $\mathbb{H}(Z)=\mathbb{H}(X)+\mathbb{H}(Y)$.

### 23.2.2.5 Proof

In the following, summation are over all possible values that the variables can have. By the independence of $X$ and $Y$ we have

$$
\begin{aligned}
\mathbb{H}(Z)= & \sum_{x, y} \operatorname{Pr}[(X, Y)=(x, y)] \lg \frac{1}{\operatorname{Pr}[(X, Y)=(x, y)]} \\
= & \sum_{x, y} \operatorname{Pr}[X=x] \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[X=x] \operatorname{Pr}[Y=y]} \\
= & \sum_{x} \sum_{y} \operatorname{Pr}[X=x] \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[X=x]} \\
& \quad+\sum_{y} \sum_{x} \operatorname{Pr}[X=x] \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[Y=y]}
\end{aligned}
$$

### 23.2.2.6 Proof continued

$$
\begin{aligned}
\mathbb{H}(Z)= & \sum_{x} \sum_{y} \operatorname{Pr}[X=x] \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[X=x]} \\
& +\sum_{y} \sum_{x} \operatorname{Pr}[X=x] \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[Y=y]} \\
= & \sum_{x} \operatorname{Pr}[X=x] \lg \frac{1}{\operatorname{Pr}[X=x]} \\
& \quad+\sum_{y} \operatorname{Pr}[Y=y] \lg \frac{1}{\operatorname{Pr}[Y=y]} \\
= & \mathbb{H}(X)+\mathbb{H}(Y) .
\end{aligned}
$$

### 23.2.2.7 Bounding the binomial coefficient using entropy

### 23.2.2.8 Bounding the binomial coefficient using entropy

Lemma 23.2.5. $q \in[0,1]$
$n q$ is integer in the range $[0, n]$.
Then

$$
\frac{2^{n \mathbb{H}(q)}}{n+1} \leq\binom{ n}{n q} \leq 2^{n \mathbb{H}(q)}
$$

### 23.2.2.9 Proof

Holds if $q=0$ or $q=1$, so assume $0<q<1$. We have

$$
\binom{n}{n q} q^{n q}(1-q)^{n-n q} \leq(q+(1-q))^{n}=1
$$

We also have: $q^{-n q}(1-q)^{-(1-q) n}=2^{n(-q \lg q-(1-q) \lg (1-q))}=2^{n \mathbb{H}(q)}$, we have

$$
\binom{n}{n q} \leq q^{-n q}(1-q)^{-(1-q) n}=2^{n \mathbb{H}(q)}
$$

### 23.2.3 Proof continued

### 23.2.3.1 Other direction...

(A) $\mu(k)=\binom{n}{k} q^{k}(1-q)^{n-k}$
(B) $\sum_{i=0}^{n}\binom{n}{i} q^{i}(1-q)^{n-i}=\sum_{i=0}^{n} \mu(i)$.
(C) Claim: $\mu(n q)=\binom{n}{n q} q^{n q}(1-q)^{n-n q}$ largest term in $\sum_{k=0}^{n} \mu(k)=1$.
(D) $\Delta_{k}=\mu(k)-\mu(k+1)=\binom{n}{k} q^{k}(1-q)^{n-k}\left(1-\frac{n-k}{k+1} \frac{q}{1-q}\right)$,
(E) sign of $\Delta_{k}=$ size of last term...
(F) $\operatorname{sign}\left(\Delta_{k}\right)=\operatorname{sign}\left(1-\frac{(n-k) q}{(k+1)(1-q)}\right)$

$$
=\operatorname{sign}\left(\frac{(k+1)(1-q)-(n-k) q}{(k+1)(1-q)}\right) .
$$

### 23.2.3.2 Proof continued

(A) $(k+1)(1-q)-(n-k) q=k+1-k q-q-n q+k q=1+k-q-n q$.
(B) $\Longrightarrow \Delta_{k} \geq 0$ when $k \geq n q+q-1$
$\Delta_{k}<0$ otherwise.
(C) $\mu(k)=\binom{n}{k} q^{k}(1-q)^{n-k}$
(D) $\mu(k)<\mu(k+1)$, for $k<n q$, and $\mu(k) \geq \mu(k+1)$ for $k \geq n q$.
(E) $\Longrightarrow \mu(n q)$ is the largest term in $\sum_{k=0}^{n} \mu(k)=1$.
(F) $\mu(n q)$ larger than the average in sum.
$(\mathrm{G}) \Longrightarrow\binom{n}{k} q^{k}(1-q)^{n-k} \geq \frac{1}{n+1}$.
$(\mathrm{H}) \Longrightarrow\binom{n}{n q} \geq \frac{1}{n+1} q^{-n q}(1-q)^{-(n-n q)}=\frac{1}{n+1} 2^{n \mathbb{H}(q)}$.

### 23.2.3.3 Generalization...

Corollary 23.2.6. We have:
(i) $q \in[0,1 / 2] \Rightarrow\binom{n}{\lfloor n \downarrow\rfloor} \leq 2^{n \mathbb{H}(q)}$. (ii) $q \in[1 / 2,1]\binom{n}{[n q\rceil} \leq 2^{n \mathbb{H}(q)}$.
(iii) $q \in[1 / 2,1] \Rightarrow \frac{2^{n \sharp(q)}}{n+1} \leq\binom{ n}{\lfloor n q\rfloor}$. (iv) $q \in[0,1 / 2] \Rightarrow \frac{2^{n \mathbb{H}(q)}}{n+1} \leq\binom{ n}{\lceil n q\rceil}$.

Proof is straightforward but tedious.

### 23.2.3.4 What we have...

(A) Proved that $\binom{n}{n q} \approx 2^{n \mathbb{H}(q)}$.
(B) Estimate is loose.
(C) Sanity check...
(I) A sequence of $n$ bits generated by coin with probability $q$ for head.
(II) By Chernoff inequality... roughly $n q$ heads in this sequence.
(III) Generated sequence $Y$ belongs to $\binom{n}{n q} \approx 2^{n \mathbb{H}(q)}$ possible sequences .
(IV) ...of similar probability.

$$
(\mathrm{V}) \Longrightarrow \mathbb{H}(Y)=n \mathbb{H}(q) \approx \lg \binom{n}{n q}
$$

### 23.2.4 Extracting randomness

### 23.2.4.1 Just one bit...

question Given a coin $C$ with:
$p$ : Probability for head.
$q=1-p$ : Probability for tail.
Q: How to get one true random bit, by flipping $C$.
Describe an algorithm!

### 23.2.4.2 Extracting randomness...

Entropy can be interpreted as the amount of unbiased random coin flips can be extracted from a random variable.

Definition 23.2.7. An extraction function Ext takes as input the value of a random variable $X$ and outputs a sequence of bits $y$, such that $\operatorname{Pr}[\operatorname{Ext}(X)=y| | y \mid=k]=\frac{1}{2^{k}}$, whenever $\operatorname{Pr}[|y|=k]>0$, where $|y|$ denotes the length of $y$.

### 23.2.4.3 Extracting randomness...

(A) $X$ : uniform random integer variable out of $0, \ldots, 7$.
(B) $\operatorname{Ext}(X)$ : binary representation of $x$.
(C) Def. subtle: all extracted seqs of same len have same probability.
(D) Another example of extraction scheme:
(A) $X$ : uniform random integer variable $0, \ldots, 11$.
(B) $\operatorname{Ext}(x)$ : output the binary representation for $x$ if $0 \leq x \leq 7$.
(C) If $x$ is between 8 and 11 ?
(D) Idea... Output binary representation of $x-8$ as a two bit number.
(E) A valid extractor...
$\operatorname{Pr}[\operatorname{Ext}(X)=00| | \operatorname{Ext}(X) \mid=2]=\frac{1}{4}$,

### 23.2.4.4 Technical lemma

The following is obvious, but we provide a proof anyway.
Lemma 23.2.8. Let $x / y$ be a faction, such that $x / y<1$. Then, for any $i$, we have $x / y<(x+i) /(y+i)$.

Proof: We need to prove that $x(y+i)-(x+i) y<0$. The left size is equal to $i(x-y)$, but since $y>x$ (as $x / y<1$ ), this quantity is negative, as required.

### 23.2.4.5 A uniform variable extractor...

Theorem 23.2.9. (A) $X$ : random variable chosen uniformly at random from $\{0, \ldots, m-1\}$.
(B) Then there is an extraction function for $X$ :
(A) outputs on average at least

$$
\lfloor\lg m\rfloor-1=\lfloor\mathbb{H}(X)\rfloor-1
$$

independent and unbiased bits.

### 23.2.4.6 Proof

(A) $m$ : A sum of unique powers of 2 , namely $m=\sum_{i} a_{i} 2^{i}$, where $a_{i} \in\{0,1\}$.
(B) Example:

(C) decomposed $\{0, \ldots, m-1\}$ into disjoint union of blocks sizes are powers of 2 .
(D) If $x$ is in block $2^{k}$, output its relative location in the block in binary representation.
(E) Example: $x=10$ :

then falls into block $2^{2} \ldots$
$x$ relative location is 2 . Output 2 written using two bits, Output: "10".

### 23.2.4.7 Proof continued

(A) Valid extractor...
(B) Theorem holds if $m$ is a power of two. Only one block.
(C) $m$ not a power of $2 \ldots$
(D) $X$ falls in block of size $2^{k}$ : then output $k$ complete random bits..
... entropy is $k$.
(E) Let $2^{k}<m<2^{k+1}$ biggest block.
(F) $u=\left\lfloor\lg \left(m-2^{k}\right)\right\rfloor<k$.

There must be a block of size $u$ in the decomposition of $m$.
(G) two blocks in decomposition of $m$ : sizes $2^{k}$ and $2^{u}$.
(H) Largest two blocks...
(I) $2^{k}+2 * 2^{u}>m \Longrightarrow 2^{u+1}+2^{k}-m>0$.
(J) $Y$ : random variable $=$ number of bits output by extractor.

### 23.2.4.8 Proof continued

(A) By lemma, since $\frac{m-2^{k}}{m}<1$ :

$$
\frac{m-2^{k}}{m} \leq \frac{m-2^{k}+\left(2^{u+1}+2^{k}-m\right)}{m+\left(2^{u+1}+2^{k}-m\right)}=\frac{2^{u+1}}{2^{u+1}+2^{k}}
$$

(B) By induction (assumed holds for all numbers smaller than $m$ ):

$$
\begin{aligned}
& \mathbf{E}[Y] \geq \frac{2^{k}}{m} k+\frac{m-2^{k}}{m}(\underbrace{\left\lfloor\lg \left(m-2^{k}\right)\right\rfloor}_{u}-1) \\
& \quad=\frac{2^{k}}{m} k+\frac{m-2^{k}}{m}(\underbrace{k-k}_{=0}+u-1) \\
& \quad=k+\frac{m-2^{k}}{m}(u-k-1)
\end{aligned}
$$

### 23.2.4.9 Proof continued..

(A) We have:

$$
\begin{aligned}
\mathbf{E}[Y] \geq k+\frac{m-2^{k}}{m} & (u-k-1) \\
\geq & k+\frac{2^{u+1}}{2^{u+1}+2^{k}}(u-k-1) \\
= & k-\frac{2^{u+1}}{2^{u+1}+2^{k}}(1+k-u)
\end{aligned}
$$

since $u-k-1 \leq 0$ as $k>u$.
(B) If $u=k-1$, then $\mathbf{E}[Y] \geq k-\frac{1}{2} \cdot 2=k-1$, as required.
(C) If $u=k-2$ then $\mathbf{E}[Y] \geq k-\frac{1}{3} \cdot 3=k-1$.

### 23.2.4.10 Proof continued.....

(A) $\mathbf{E}[Y] \geq k-\frac{2^{u+1}}{2^{u+1}+2^{k}}(1+k-u)$. And $u-k-1 \leq 0$ as $k>u$.
(B) If $u<k-2$ then

$$
\begin{aligned}
\mathbf{E}[Y] & \geq k-\frac{2^{u+1}}{2^{k}}(1+k-u) \\
& =k-\frac{k-u+1}{2^{k-u-1}} \\
& =k-\frac{2+(k-u-1)}{2^{k-u-1}} \\
& \geq k-1
\end{aligned}
$$

since $(2+i) / 2^{i} \leq 1$ for $i \geq 2$.

