## Chapter 23

## Entropy, Randomness, and Information

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## 23.1 Entropy

#### 23.1.0.1 Quote

"If only once - only once - no matter where, no matter before what audience - I could better the record of the great Rastelli and juggle with thirteen balls, instead of my usual twelve, I would feel that I had truly accomplished something for my country. But I am not getting any younger, and although I am still at the peak of my powers there are moments - why deny it? - when I begin to doubt - and there is a time limit on all of us."

-Romain Gary, The talent scout.

## 23.2 Entropy

#### 23.2.0.2 Entropy: Definition

Definition 23.2.1. The *entropy* in bits of a discrete random variable X is

$$\mathbb{H}(X) = -\sum_{x} \mathbf{Pr} \left[ X = x \right] \lg \mathbf{Pr} \left[ X = x \right].$$

Equivalently,  $\mathbb{H}(X) = \mathbf{E}\left[\lg \frac{1}{\mathbf{Pr}[X]}\right].$ 

#### 23.2.0.3 Entropy intuition...

Intuition...  $\mathbb{H}(X)$  is the number of **fair** coin flips that one gets when getting the value of X.

Interpretation from last lecture... Consider a (huge) string  $S = s_1 s_2 \dots s_n$  formed by picking characters independently according to X. Then

$$|S| \mathbb{H}(X) = n \mathbb{H}(X)$$

is the minimum number of bits one needs to store the string S.

#### 23.2.0.4 Binary entropy

$$\mathbb{H}(X) = -\sum_{x} \mathbf{Pr} \Big[ X = x \Big] \lg \mathbf{Pr} \Big[ X = x \Big]$$

**Definition 23.2.2.** The **binary entropy** function  $\mathbb{H}(p)$  for a random binary variable that is 1 with probability p, is  $\mathbb{H}(p) = -p \lg p - (1-p) \lg (1-p)$ . We define  $\mathbb{H}(0) = \mathbb{H}(1) = 0$ .

Q: How many truly random bits are there when given the result of flipping a single coin with probability p for heads?



**23.2.0.5** Binary entropy:  $\mathbb{H}(p) = -p \lg p - (1-p) \lg (1-p)$ 

- (A)  $\mathbb{H}(p)$  is a concave symmetric around 1/2 on the interval [0, 1].
- (B) maximum at 1/2.
- (C)  $\mathbb{H}(3/4) \approx 0.8113$  and  $\mathbb{H}(7/8) \approx 0.5436$ .
- (D)  $\implies$  coin that has 3/4 probably to be heads have higher amount of "randomness" in it than a coin that has probability 7/8 for heads.

#### 23.2.0.6 And now for some unnecessary math

(A)  $\mathbb{H}(p) = -p \lg p - (1-p) \lg (1-p)$ 

(B) 
$$\mathbb{H}'(p) = -\lg p + \lg(1-p) = \lg \frac{1-p}{p}$$

- (C)  $\mathbb{H}''(p) = \frac{p}{1-p} \cdot \left(-\frac{1}{p^2}\right) = -\frac{1}{p(1-p)}.$
- (D)  $\implies \mathbb{H}''(p) \leq 0$ , for all  $p \in (0, 1)$ , and the  $\mathbb{H}(\cdot)$  is concave.
- (E)  $\mathbb{H}'(1/2) = 0 \implies \mathbb{H}(1/2) = 1$  max of binary entropy.

(F)  $\implies$  balanced coin has the largest amount of randomness in it.

#### 23.2.1 Task at hand: Squeezing good random bits...

#### 23.2.1.1 ...out of bad random bits...

- (A)  $b_1, \ldots, b_n$ : result of *n* coin flips...
- (B) From a faulty coin!
- (C) p: probability for head.
- (D) We need fair bit coins!
- (E) Convert  $b_1, \ldots, b_n \implies b'_1, \ldots, b'_m$ .
- (F) New bits must be truly random: Probability for head is 1/2.
- (G)  $\mathbf{Q}$ : How many truly random bits can we extract?

### 23.2.2 Intuitively...

#### 23.2.2.1 Squeezing good random bits out of bad random bits...

Question... Given the result of n coin flips:  $b_1, \ldots, b_n$  from a faulty coin, with head with probability p, how many truly random bits can we extract?

If believe intuition about entropy, then this number should be  $\approx n\mathbb{H}(p)$ .

#### 23.2.2.2 Back to Entropy

- (A) *entropy* of X is  $\mathbb{H}(X) = -\sum_{x} \Pr[X = x] \lg \Pr[X = x].$
- (B) Entropy of uniform variable.

**Example 23.2.3.** A random variable X that has probability 1/n to be *i*, for i = 1, ..., n, has entropy  $\mathbb{H}(X) = -\sum_{i=1}^{n} \frac{1}{n} \lg \frac{1}{n} = \lg n$ .

- (C) Entropy is oblivious to the exact values random variable can have.
- (D)  $\implies$  random variables over -1, +1 with equal probability has the same entropy (i.e., 1) as a fair coin.

#### 23.2.2.3 Lemma: Entropy additive for independent variables

#### 23.2.2.4 Lemma: Entropy additive for independent variables

**Lemma 23.2.4.** Let X and Y be two independent random variables, and let Z be the random variable (X, Y). Then  $\mathbb{H}(Z) = \mathbb{H}(X) + \mathbb{H}(Y)$ .

#### 23.2.2.5 Proof

In the following, summation are over all possible values that the variables can have. By the independence of X and Y we have

$$\mathbb{H}(Z) = \sum_{x,y} \mathbf{Pr}\Big[(X,Y) = (x,y)\Big] \lg \frac{1}{\mathbf{Pr}[(X,Y) = (x,y)]}$$
$$= \sum_{x,y} \mathbf{Pr}\Big[X = x\Big] \mathbf{Pr}\Big[Y = y\Big] \lg \frac{1}{\mathbf{Pr}[X = x] \mathbf{Pr}[Y = y]}$$
$$= \sum_{x} \sum_{y} \mathbf{Pr}[X = x] \mathbf{Pr}[Y = y] \lg \frac{1}{\mathbf{Pr}[X = x]}$$
$$+ \sum_{y} \sum_{x} \mathbf{Pr}[X = x] \mathbf{Pr}[Y = y] \lg \frac{1}{\mathbf{Pr}[Y = y]}$$

#### 23.2.2.6 Proof continued

$$\begin{split} \mathbb{H}(Z) &= \sum_{x} \sum_{y} \mathbf{Pr}[X=x] \, \mathbf{Pr}[Y=y] \lg \frac{1}{\mathbf{Pr}[X=x]} \\ &+ \sum_{y} \sum_{x} \mathbf{Pr}[X=x] \, \mathbf{Pr}[Y=y] \lg \frac{1}{\mathbf{Pr}[Y=y]} \\ &= \sum_{x} \mathbf{Pr}[X=x] \lg \frac{1}{\mathbf{Pr}[X=x]} \\ &+ \sum_{y} \mathbf{Pr}[Y=y] \lg \frac{1}{\mathbf{Pr}[Y=y]} \\ &= \mathbb{H}(X) + \mathbb{H}(Y) \, . \end{split}$$

# 23.2.2.7 Bounding the binomial coefficient using entropy23.2.2.8 Bounding the binomial coefficient using entropy

#### Lemma 23.2.5. $q \in [0, 1]$

nq is integer in the range [0, n]. Then

$$\frac{2^{n\mathbb{H}(q)}}{n+1} \le \binom{n}{nq} \le 2^{n\mathbb{H}(q)}.$$

#### 23.2.2.9 Proof

Holds if q = 0 or q = 1, so assume 0 < q < 1. We have

$$\binom{n}{nq}q^{nq}(1-q)^{n-nq} \le (q+(1-q))^n = 1.$$

We also have:  $q^{-nq}(1-q)^{-(1-q)n} = 2^{n(-q \lg q - (1-q) \lg (1-q))} = 2^{n \mathbb{H}(q)}$ , we have

$$\binom{n}{nq} \le q^{-nq}(1-q)^{-(1-q)n} = 2^{n\mathbb{H}(q)}.$$

### 23.2.3 Proof continued

#### 23.2.3.1 Other direction...

$$\begin{aligned} \text{(A)} \quad \mu(k) &= \binom{n}{k} q^k (1-q)^{n-k} \\ \text{(B)} \quad \sum_{i=0}^n \binom{n}{i} q^i (1-q)^{n-i} &= \sum_{i=0}^n \mu(i). \\ \text{(C)} \quad \text{Claim:} \quad \mu(nq) &= \binom{n}{nq} q^{nq} (1-q)^{n-nq} \text{ largest term in } \sum_{k=0}^n \mu(k) = 1. \\ \text{(D)} \quad \Delta_k &= \mu(k) - \mu(k+1) = \binom{n}{k} q^k (1-q)^{n-k} \left(1 - \frac{n-k}{k+1} \frac{q}{1-q}\right), \\ \text{(E)} \quad \text{sign of } \Delta_k &= \text{size of last term...} \\ \text{(F)} \quad \text{sign}(\Delta_k) &= \text{sign} \left(1 - \frac{(n-k)q}{(k+1)(1-q)}\right) \\ &= \text{sign} \left(\frac{(k+1)(1-q)-(n-k)q}{(k+1)(1-q)}\right). \end{aligned}$$

#### 23.2.3.2 Proof continued

 $\begin{array}{ll} (\mathrm{A}) & (k+1)(1-q) - (n-k)q = k+1 - kq - q - nq + kq = 1 + k - q - nq. \\ (\mathrm{B}) & \Longrightarrow & \Delta_k \geq 0 \text{ when } k \geq nq + q - 1 \\ & \Delta_k < 0 \text{ otherwise.} \\ (\mathrm{C}) & \mu(k) = \binom{n}{k} q^k (1-q)^{n-k} \\ (\mathrm{D}) & \mu(k) < \mu(k+1), \text{ for } k < nq, \text{ and } \mu(k) \geq \mu(k+1) \text{ for } k \geq nq. \\ (\mathrm{E}) & \Longrightarrow & \mu(nq) \text{ is the largest term in } \sum_{k=0}^n \mu(k) = 1. \\ (\mathrm{F}) & \mu(nq) \text{ larger than the average in sum.} \\ (\mathrm{G}) & \Longrightarrow & \binom{n}{k} q^k (1-q)^{n-k} \geq \frac{1}{n+1}. \end{array}$ 

(H) 
$$\implies (n_{nq}) \ge \frac{1}{n+1}q^{-nq}(1-q)^{-(n-nq)} = \frac{1}{n+1}2^{n\mathbb{H}(q)}.$$

#### 23.2.3.3 Generalization...

**Corollary 23.2.6.** We have:  
(i) 
$$q \in [0, 1/2] \Rightarrow \binom{n}{\lfloor nq \rfloor} \leq 2^{n\mathbb{H}(q)}$$
. (ii)  $q \in [1/2, 1] \binom{n}{\lceil nq \rceil} \leq 2^{n\mathbb{H}(q)}$ .  
(iii)  $q \in [1/2, 1] \Rightarrow \frac{2^{n\mathbb{H}(q)}}{n+1} \leq \binom{n}{\lfloor nq \rfloor}$ . (iv)  $q \in [0, 1/2] \Rightarrow \frac{2^{n\mathbb{H}(q)}}{n+1} \leq \binom{n}{\lceil nq \rceil}$ 

Proof is straightforward but tedious.

#### 23.2.3.4 What we have...

- (A) Proved that  $\binom{n}{nq} \approx 2^{n\mathbb{H}(q)}$ .
- (B) Estimate is loose.
- (C) Sanity check...
  - (I) A sequence of n bits generated by coin with probability q for head.
  - (II) By Chernoff inequality... roughly nq heads in this sequence.
  - (III) Generated sequence Y belongs to  $\binom{n}{nq} \approx 2^{n\mathbb{H}(q)}$  possible sequences.
  - (IV) ... of similar probability.
  - (V)  $\implies \mathbb{H}(Y) = n\mathbb{H}(q) \approx \lg \binom{n}{nq}.$

### 23.2.4 Extracting randomness

#### 23.2.4.1 Just one bit...

question Given a coin C with:

- p: Probability for head.
- q = 1 p: Probability for tail.
- **Q:** How to get <u>one</u> true random bit, by flipping C. Describe an algorithm!

#### 23.2.4.2 Extracting randomness...

Entropy can be interpreted as the amount of unbiased random coin flips can be extracted from a random variable.

**Definition 23.2.7.** An extraction function **Ext** takes as input the value of a random variable X and outputs a sequence of bits y, such that  $\mathbf{Pr}\left[\mathbf{Ext}(X) = y \mid |y| = k\right] = \frac{1}{2^k}$ , whenever  $\mathbf{Pr}[|y| = k] > 0$ , where |y| denotes the length of y.

#### 23.2.4.3 Extracting randomness...

- (A) X: uniform random integer variable out of  $0, \ldots, 7$ .
- (B)  $\mathbf{Ext}(X)$ : binary representation of x.
- (C) Def. subtle: all extracted seqs of same len have same probability.
- (D) Another example of extraction scheme:
  - (A) X: uniform random integer variable  $0, \ldots, 11$ .
  - (B)  $\mathbf{Ext}(x)$ : output the binary representation for x if  $0 \le x \le 7$ .
  - (C) If x is between 8 and 11?
  - (D) Idea... Output binary representation of x 8 as a two bit number.
- (E) A valid extractor...

 $\mathbf{Pr}\Big[\mathbf{Ext}(X) = 00 \ \Big| \ |\mathbf{Ext}(X)| = 2\Big] = \frac{1}{4},$ 

#### 23.2.4.4 Technical lemma

The following is obvious, but we provide a proof anyway.

**Lemma 23.2.8.** Let x/y be a faction, such that x/y < 1. Then, for any i, we have x/y < (x+i)/(y+i).

*Proof:* We need to prove that x(y+i) - (x+i)y < 0. The left size is equal to i(x-y), but since y > x (as x/y < 1), this quantity is negative, as required.

#### 23.2.4.5 A uniform variable extractor...

**Theorem 23.2.9.** (A) X: random variable chosen uniformly at random from  $\{0, ..., m-1\}$ . (B) Then there is an extraction function for X:

(A) outputs on average at least

$$\left\lfloor \lg m \right\rfloor - 1 = \left\lfloor \mathbb{H}\left(X\right) \right\rfloor - 1$$

independent and unbiased bits.

#### 23.2.4.6 Proof

(A) m: A sum of unique powers of 2, namely  $m = \sum_i a_i 2^i$ , where  $a_i \in \{0, 1\}$ .

	0123456	<b>5789101214</b>	$\overline{0\ 1\ 2\ 3\ 4\ 5\ 6\ 7\ 8}$	$9\ 10\ 12\ 14$
(B) Example:		11 13		11 13

- (C) decomposed  $\{0, \ldots, m-1\}$  into disjoint union of blocks sizes are powers of 2.
- (D) If x is in block  $2^k$ , output its relative location in the block in binary representation.
- (E) Example: x = 10: then falls into block  $2^2$ ... x relative location is 2. Output 2 written using two bits, Output: "10".

#### 23.2.4.7Proof continued

- (A) Valid extractor...
- (B) Theorem holds if m is a power of two. Only one block.
- (C) m not a power of 2...
- (D) X falls in block of size  $2^k$ : then output k complete random bits..  $\dots$  entropy is k.
- (E) Let  $2^k < m < 2^{k+1}$  biggest block.
- (F)  $u = \lfloor \lg(m-2^k) \rfloor < k$ . There must be a block of size u in the decomposition of m.
- (G) two blocks in decomposition of m: sizes  $2^k$  and  $2^u$ .
- (H) Largest two blocks...
- (I)  $2^k + 2 * 2^u > m \implies 2^{u+1} + 2^k m > 0.$
- (J) Y: random variable = number of bits output by extractor.

#### 23.2.4.8 Proof continued

(A) By lemma, since  $\frac{m-2^k}{m} < 1$ :

$$\frac{m-2^k}{m} \le \frac{m-2^k + \left(2^{u+1} + 2^k - m\right)}{m + \left(2^{u+1} + 2^k - m\right)} = \frac{2^{u+1}}{2^{u+1} + 2^k}$$

(B) By induction (assumed holds for all numbers smaller than m):

$$\mathbf{E}[Y] \ge \frac{2^k}{m}k + \frac{m-2^k}{m} \left( \underbrace{\lfloor \lg(m-2^k) \rfloor}_u - 1 \right)$$
$$= \frac{2^k}{m}k + \frac{m-2^k}{m} \underbrace{(k-k)}_{=0} + u - 1$$

$$=k+\frac{m-2^{k}}{m}(u-k-1)$$

#### 23.2.4.9 Proof continued..

(A) We have:

$$\begin{split} \mathbf{E}\Big[Y\Big] &\geq k + \frac{m-2^k}{m}(u-k-1) \\ &\geq k + \frac{2^{u+1}}{2^{u+1}+2^k} \left(u-k-1\right) \\ &= k - \frac{2^{u+1}}{2^{u+1}+2^k} (1+k-u) \,, \end{split}$$

since  $u - k - 1 \leq 0$  as k > u.

- (B) If u = k 1, then  $\mathbf{E}[Y] \ge k \frac{1}{2} \cdot 2 = k 1$ , as required. (C) If u = k 2 then  $\mathbf{E}[Y] \ge k \frac{1}{3} \cdot 3 = k 1$ .

## 23.2.4.10 Proof continued.....

(A) 
$$\mathbf{E}[Y] \ge k - \frac{2^{u+1}}{2^{u+1}+2^k}(1+k-u).$$
  
And  $u-k-1 \le 0$  as  $k > u.$ 

(B) If u < k - 2 then

$$\mathbf{E}[Y] \ge k - \frac{2^{u+1}}{2^k}(1+k-u) \\ = k - \frac{k-u+1}{2^{k-u-1}} \\ = k - \frac{2+(k-u-1)}{2^{k-u-1}} \\ \ge k-1,$$

since  $(2+i)/2^i \le 1$  for  $i \ge 2$ .