CS 473: Algorithms

Ruta Mehta

University of Illinois, Urbana-Champaign

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CS 473: Algorithms, Spring 2021

Inequalities & Randomized QuickSort

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Most slides are courtesy Prof. Chekuri

Ruta (UIUC)

CS473

Outline

Slick Analysis of Randomized QuickSort

Concentration of Mass Around Mean

Markov's Inequality

Chebyshev's Inequality

Chernoff Bound

Randomized **QuickSort**: High Probability Analysis

Part I

Analysis of QuickSort

Recall: Randomized QuickSort

Randomized QuickSort

- **1** Pick a pivot element *uniformly at random* from the array.
- Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- In the subarrays, and concatenate them.

Theorem

Expected running time of Randomized QuickSort on an array of size n is $O(n \log n)$.

- A: Given array with *n* distinct numbers.
- Q(A): number of comparisons of randomized QuickSort on A.
 Note that Q(A) is a random variable.
- **3** X_i : Random variable indicating if picked pivot has rank i in A.

 A_{left}^{i} and A_{right}^{i} be the corresponding left and right subarrays.

$$Q(A) = n + \sum_{i=1}^{n} X_i \cdot \left(Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i)\right).$$

Exactly one non-zero X_i . $E[X_i] = Pr[pivot has rank i] = 1/n$.

Independence of Random Variables

Lemma

Random variables X_i is independent of random variables $Q(A_{left}^i)$ as well as $Q(A_{right}^i)$, i.e.

$$\mathbf{E} \begin{bmatrix} X_i \cdot Q(A_{left}^i) \end{bmatrix} = \mathbf{E} \begin{bmatrix} X_i \end{bmatrix} \mathbf{E} \begin{bmatrix} Q(A_{left}^i) \end{bmatrix}$$
$$\mathbf{E} \begin{bmatrix} X_i \cdot Q(A_{right}^i) \end{bmatrix} = \mathbf{E} \begin{bmatrix} X_i \end{bmatrix} \mathbf{E} \begin{bmatrix} Q(A_{right}^i) \end{bmatrix}$$

Proof.

This is because the algorithm, while recursing on $Q(A_{left}^{i})$ and $Q(A_{right}^{i})$ uses new random coin tosses that are independent of the coin tosses used to decide the first pivot. Only the latter decides value of X_{i} .

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 $T(n) = \max_{A:|A|=n} E[Q(A)]$ be the worst-case expected running time on arrays of size n.

We have, for any **A**:

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By linearity of expectation, and independence random variables:

 $T(n) = \max_{A:|A|=n} E[Q(A)]$ be the worst-case expected running time on arrays of size n. We derived:

$$\mathsf{E}\Big[Q(A)\Big] \leq n + \sum_{i=1}^{n} \frac{1}{n} \left(T(i-1) + T(n-i)\right).$$

Note that above holds for any A of size n. Therefore

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$$T(n) = \max_{A:|A|=n} \mathbb{E}[Q(A)] \le n + \sum_{i=1}^{n} \frac{1}{n} (T(i-1) + T(n-i)).$$

Solving the Recurrence

$$T(n) \leq n + \sum_{i=1}^{n} \frac{1}{n} (T(i-1) + T(n-i))$$

with base case T(1) = 0.

Lemma

 $T(n) = O(n \log n).$

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 $T(n) = O(n \log n).$

Proof.

(Guess and) Verify by induction.

Part II

Slick analysis of QuickSort

Q(A): number of comparisons done on input array A

- Sank of an element is its position in the sorted A.
- R_{ij}: event that rank *i* element is compared with rank *j* element, for 1 ≤ *i* < *j* < *n*.

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- Sank of an element is its position in the sorted A.
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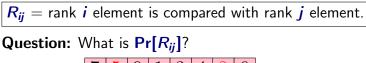
$$Q(A) = \sum_{1 \le i < j \le n} X_{ij}$$

and hence by linearity of expectation,

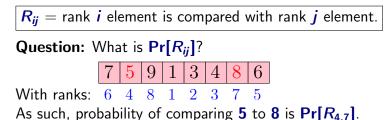
$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \le i < j \le n} \mathsf{E}\Big[X_{ij}\Big] = \sum_{1 \le i < j \le n} \mathsf{Pr}\Big[R_{ij}\Big].$$

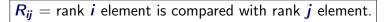
 R_{ij} = rank *i* element is compared with rank *j* element.

Question: What is Pr[R_{ij}]?



With ranks: $6 \ 4 \ 8 \ 1 \ 2 \ 3 \ 7 \ 5$





Question: What is **Pr**[*R*_{ij}]?

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If pivot too small (say 3 [rank 2]). Partition and call recursively:

Decision if to compare **5** to **8** is moved to subproblem.

5 | 9



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If pivot too large (say 9 [rank 8]):

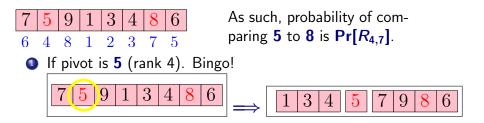
3 | 4

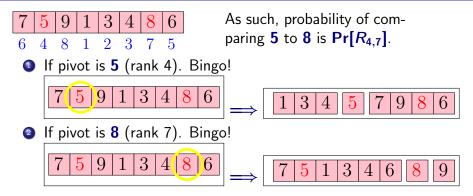
9 1 3 4 8 6

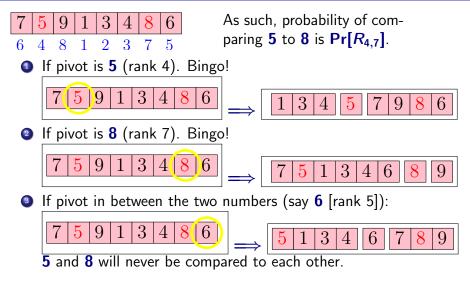
$$\overrightarrow{5} \implies \boxed{75134869}$$

Decision if to compare 5 to 8 moved to subproblem.

5







Conclusion:

R_{*i*,*j*} happens if and only if:

*i*th or *j*th ranked element is the first pivot out of *i*th to *j*th ranked elements.

 $\Pr[R_{i,j}] = \Pr[i \text{th or } j \text{th ranked element is the pivot } | \\pivot has rank in \{i, i + 1, \dots, j - 1, j\}]$

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There are k = j - i + 1 relevant elements.

$$\Pr\left[R_{i,j}\right] = \frac{2}{k} = \frac{2}{j-i+1}.$$

Question: What is **Pr**[*R*_{*ij*}]?

Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

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Lemma $\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$

Proof.

Let $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$ be elements of A in sorted order. Let $S = \{a_i, a_{i+1}, \ldots, a_j\}$

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A Slick Analysis of **QuickSort** Continued...

Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

Proof.

Let $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$ be sort of A. Let $S = \{a_i, a_{i+1}, \ldots, a_j\}$ **Observation:** a_i is compared with a_j if and only if either a_i or a_j is chosen as a pivot from S at separation. **Observation:** Given that pivot is chosen from S the probability that it is a_i or a_j is exactly 2/|S| = 2/(j - i + 1) since the pivot is chosen uniformly at random from the array.

A Slick Analysis of **QuickSort** Continued...

$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \leq i < j \leq n} \mathsf{E}[X_{ij}] = \sum_{1 \leq i < j \leq n} \mathsf{Pr}[R_{ij}].$$

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A Slick Analysis of **QuickSort** Continued...

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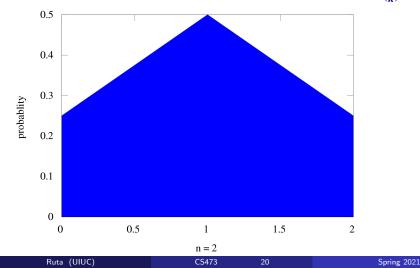
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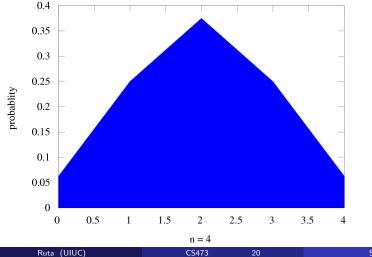
Part III

Inequalities

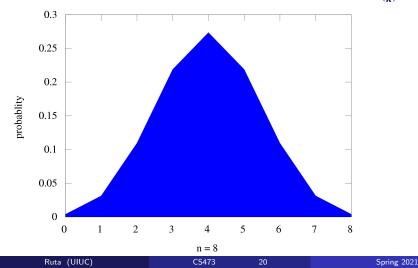
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p. $\binom{n}{k} \frac{1}{2^n}$.



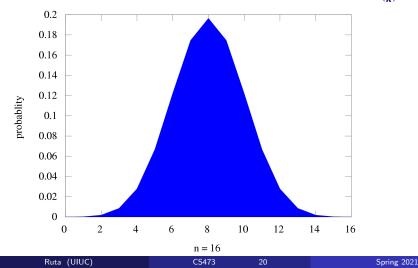
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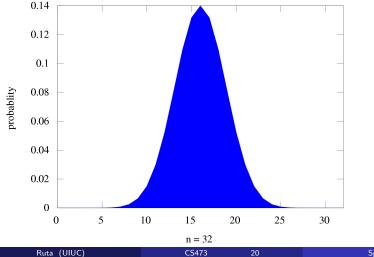
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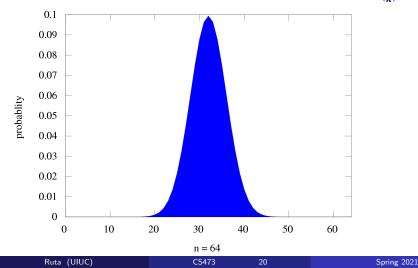
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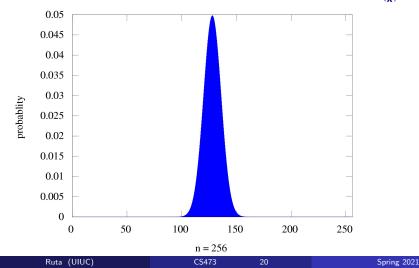
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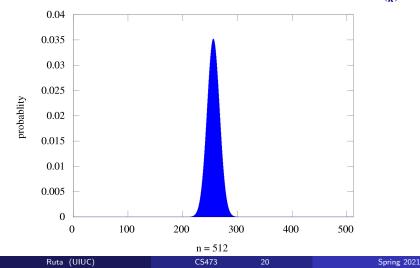
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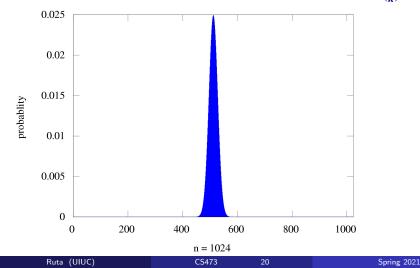
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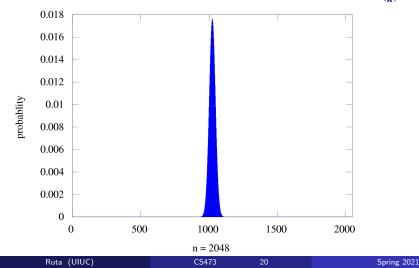
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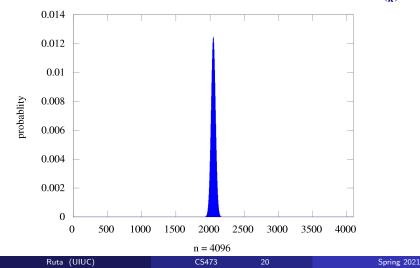
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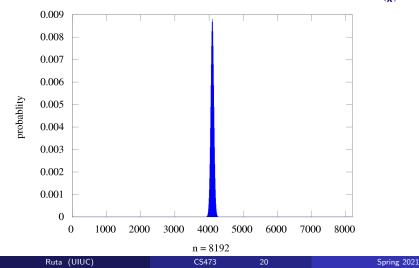
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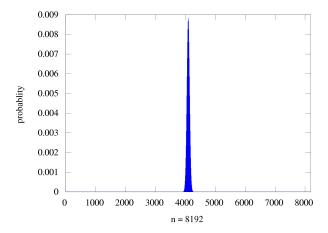


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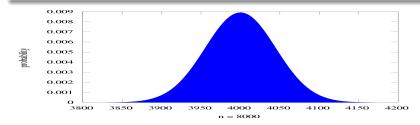




This is known as **concentration of mass**. This is a very special case of the **law of large numbers**.

Informal statement of law of large numbers

For n large enough, the middle portion of the binomial distribution looks like (converges to) the normal/Gaussian distribution.



Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

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Use of well known inequalities in analysis.

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Analysis

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- Suppose $\Pr[Q \ge 10 n lgn] \le c$. Also we know that $Q \le n^2$.
- $E[Q] \le (10n \log n)(1-c) + n^2 c$

Question:

How to find c, or in other words bound $\Pr[Q \ge 10n \log n]$?

Markov's Inequality

Markov's inequality

Let X be a **non-negative** random variable over a probability space (Ω, \Pr) . For any a > 0,

$$\Pr[X \ge a] \le \frac{\mathsf{E}[X]}{a}$$

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Proof:

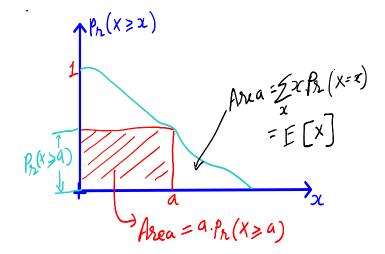
$$E[X] = \sum_{\omega \in \Omega} X(\omega) \Pr[\omega]$$

$$\geq \sum_{\omega \in \Omega, \ X(\omega) \geq a} X(\omega) \Pr[\omega]$$

$$\geq a \sum_{\omega \in \Omega, \ X(\omega) \geq a} \Pr[\omega]$$

$$= a \Pr[X \geq a]$$

Markov's Inequality: Proof by Picture



- *n* black and white balls in a bin.
- We wish to estimate the fraction of black balls. Lets say it is p^* .

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Question

How large k needs to be before our estimated value p is close to p^* ?

A rough estimate through Markov's inequality.

Lemma

For any $k \geq 1$ and p = B/k, $\Pr[p \geq 2p^*] \leq \frac{1}{2}$

A rough estimate through Markov's inequality.

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Proof.

- For each 1 ≤ i ≤ k define random variable X_i, which is 1 if ith ball is black, otherwise 0.
- $E[X_i] = Pr[X_i = 1] = p^*$.

Example: Balls in a bin

A rough estimate through Markov's inequality.

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- $E[X_i] = Pr[X_i = 1] = p^*$.
- $B = \sum_{i=1}^{k} X_i$, then $E[B] = \sum_{i=1}^{k} E[X_i] = kp^*$.

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For any
$$k\geq 1$$
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- For each 1 ≤ i ≤ k define random variable X_i, which is 1 if ith ball is black, otherwise 0.
- $E[X_i] = Pr[X_i = 1] = p^*$.
- $B = \sum_{i=1}^{k} X_i$, then $E[B] = \sum_{i=1}^{k} E[X_i] = kp^*$.
- Markov's inequality gives, $\Pr[p \ge 2p^*] =$

$$\Pr\left[\frac{B}{k} \ge 2p^*\right] = \Pr[B \ge 2kp^*] = \Pr[B \ge 2\operatorname{E}[B]] \le \frac{1}{2}$$

Variance

Given a random variable X over probability space (Ω, Pr) , variance of X is the measure of how much does it deviate from its mean value. Formally, $Var(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$

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Intuitive Derivation

Define $Y = (X - E[X])^2 = X^2 - 2X E[X] + E[X]^2$.

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Intuitive Derivation

Define
$$Y = (X - E[X])^2 = X^2 - 2X E[X] + E[X]^2$$
.

$$Var(X) = E[Y] = E[X^{2}] - 2 E[X] E[X] + E[X]^{2} = E[X^{2}] - E[X]^{2}$$

Independence

Random variables X and Y are called mutually independent if $\forall x, y \in \mathbb{R}, \ \Pr[X = x \land Y = y] = \Pr[X = x] \Pr[Y = y]$

Lemma

If X and Y are independent random variables then Var(X + Y) = Var(X) + Var(Y).

Independence

Random variables X and Y are called mutually independent if $\forall x, y \in \mathbb{R}, \ \Pr[X = x \land Y = y] = \Pr[X = x] \Pr[Y = y]$

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Chebyshev's Inequality

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 $Y = (X - E[X])^2$ is a non-negative random variable. Apply Markov's Inequality to Y for a^2 .

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$$\begin{aligned} &\mathsf{Pr}[X \leq \mathsf{E}[X] - a] \leq \frac{\operatorname{Var}(X)}{a^2} \text{ AND} \\ &\mathsf{Pr}[X \geq \mathsf{E}[X] + a] \leq \frac{\operatorname{Var}(X)}{a^2} \end{aligned}$$

Lemma

For
$$0 < \epsilon < 1$$
, $k \ge 1$ and $p = B/k$, $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$.

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• Recall: X_i is 1 if i^{th} ball is black, else 0, $B = \sum_{i=1}^k X_i$. $E[X_i] = p^*$, $E[B] = kp^*$. $p = {}^B/k$.

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$$\begin{aligned} \Pr[|p - p^*| \ge \epsilon] &= \Pr[|B/k - p^*| \ge \epsilon] \\ &= \Pr[|B - kp^*| \ge k\epsilon] \\ (\text{Chebyshev}) &\le \frac{\operatorname{Var}(B)}{k^2 \epsilon^2} = \frac{kp^*(1 - p^*)}{k^2 \epsilon^2} \\ &< \frac{1}{k\epsilon^2} \end{aligned}$$

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Let X_1, \ldots, X_k be k independent binary random variables such that, for each $i \in [1, k]$, X_i equals 1 w.p. p_i , and 0 w.p. $(1 - p_i)$.

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In notes!

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2

Example Summary

The problem was to estimate the fraction of black balls p^* in a bin filled with white and black balls. Our estimate was $p = \frac{B}{k}$ instead, where out of k draws (with replacement) B balls turns out black.

Markov's Inequality

For any $k \geq 1$, $\Pr[p \geq 2p^*] \leq \frac{1}{2}$

Chebyshev's Inequality

For any $0 < \epsilon < 1$, and $k \ge 1$, $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$.

Chernoff Bound

For any $0 < \epsilon < 1$, and $k \ge 1$, $\Pr[|p - p^*| \ge \epsilon] \le 2e^{-\frac{k\epsilon^2}{3}}$.

Part IV

Randomized QuickSort (Contd.)

Randomized QuickSort: Recall

Input: Array A of n numbers. Output: Numbers in sorted order.

Randomized QuickSort

- **1** Pick a pivot element *uniformly at random* from **A**.
- Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- In the subarrays, and concatenate them.

Note: On *every* input randomized **QuickSort** takes $O(n \log n)$ time in expectation. On *every* input it may take $\Omega(n^2)$ time with some small probability.

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- Secursively sort the subarrays, and concatenate them.

Note: On *every* input randomized **QuickSort** takes $O(n \log n)$ time in expectation. On *every* input it may take $\Omega(n^2)$ time with some small probability.

Question: With what probability it takes $O(n \log n)$ time?

Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that $\Pr[Q(A) \leq 32n \ln n] \geq 1 - 1/n^3$.

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If n = 100 then this gives $\Pr[Q(A) \le 32n \ln n] \ge 0.999999$.

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Outline of the proof

• k: depth of the recursion. Then $Q(A) \leq kn$.

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 - 3 Therefore, all elements participate in $\leq 32 \ln n$ w.p. $(1 1/n^3)$.

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- If $\rho = \#$ lucky rounds in first k rounds, then $|S_k| \leq (3/4)^{\rho} n$.
- For $|S_k| = 1$, $\rho = \log_{4/3} n \le 4 \ln n$ suffices.

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$$\begin{aligned} \Pr[\rho \le 4 \ln n] &= \Pr[\rho \le \frac{k}{8}] \\ &= \Pr[\rho \le (1 - \delta)\mu] \\ (Chernoff) &\le 2e^{\frac{-\delta^2 \mu}{2}} = 2e^{-\frac{9k}{64}} \\ &= 2e^{-4.5 \ln n} \le \frac{1}{n^4} \end{aligned}$$

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With high probability (i.e., $1 - \frac{1}{n^3}$) the depth of the recursion of **QuickSort** is $\leq 32 \ln n$. Due to *n* comparisons in each level, with high probability, the running time of **QuickSort** is $O(n \ln n)$.

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Q: How to increase the probability?