# CS 473: Algorithms 

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## CS 473: Algorithms, Spring 2021

## Inequalities \& Randomized QuickSort

Lecture 8
Feb 18, 2021

Most slides are courtesy Prof. Chekuri

## Outline

## Slick Analysis of Randomized QuickSort

## Concentration of Mass Around Mean

Markov's Inequality

Chebyshev's Inequality

Chernoff Bound

Randomized QuickSort: High Probability Analysis

## Part I

## Analysis of QuickSort

## Recall: Randomized QuickSort

## Randomized QuickSort

(1) Pick a pivot element uniformly at random from the array.
(2) Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
(3) Recursively sort the subarrays, and concatenate them.

## Theorem

Expected running time of Randomized QuickSort on an array of size $n$ is $O(n \log n)$.

## Analysis via Recurrence

(1) A: Given array with $n$ distinct numbers.
(2) $Q(A)$ : number of comparisons of randomized QuickSort on $A$. Note that $Q(A)$ is a random variable.
(3) $X_{i}$ : Random variable indicating if picked pivot has rank $\boldsymbol{i}$ in $\boldsymbol{A}$. $\boldsymbol{A}_{\text {left }}^{i}$ and $\boldsymbol{A}_{\text {right }}^{i}$ be the corresponding left and right subarrays.

$$
Q(A)=n+\sum_{i=1}^{n} X_{i} \cdot\left(Q\left(A_{\mathrm{left}}^{i}\right)+Q\left(A_{\mathrm{right}}^{i}\right)\right)
$$

Exactly one non-zero $X_{i} . E\left[X_{i}\right]=\operatorname{Pr}[$ pivot has rank $i]=1 / n$.

## Independence of Random Variables

## Lemma

Random variables $\boldsymbol{X}_{\boldsymbol{i}}$ is independent of random variables $\mathbf{Q}\left(\boldsymbol{A}_{\text {left }}^{i}\right)$ as well as $Q\left(A_{\text {right }}^{i}\right)$, i.e.

$$
\begin{aligned}
\mathrm{E}\left[X_{i} \cdot Q\left(A_{\text {left }}^{i}\right)\right] & =\mathrm{E}\left[X_{i}\right] \mathrm{E}\left[Q\left(A_{\text {left }}^{i}\right)\right] \\
\mathrm{E}\left[X_{i} \cdot Q\left(A_{\text {right }}^{i}\right)\right] & =\mathrm{E}\left[X_{i}\right] \mathrm{E}\left[Q\left(A_{\text {right }}^{i}\right)\right]
\end{aligned}
$$

## Proof.

This is because the algorithm, while recursing on $Q\left(A_{\text {left }}^{i}\right)$ and $Q\left(A_{\text {right }}^{i}\right)$ uses new random coin tosses that are independent of the coin tosses used to decide the first pivot. Only the latter decides value of $\boldsymbol{X}_{\boldsymbol{i}}$.

## Analysis via Recurrence

$T(n)=\max _{A:|A|=\boldsymbol{n}} E[Q(A)]$ be the worst-case expected running time on arrays of size $\boldsymbol{n}$.

We have, for any $\boldsymbol{A}$ :

$$
Q(A)=n+\sum_{i=1}^{n} X_{i}\left(Q\left(A_{\mathrm{left}}^{i}\right)+Q\left(A_{\mathrm{right}}^{i}\right)\right)
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$$

By linearity of expectation, and independence random variables:

$$
\begin{aligned}
\mathrm{E}[Q(A)] & =n+\sum_{i=1}^{n} \mathrm{E}\left[X_{i}\right]\left(\mathrm{E}\left[Q\left(A_{\text {left }}^{i}\right)\right]+\mathrm{E}\left[Q\left(A_{\text {right }}^{i}\right)\right]\right) \\
& \leq n+\sum_{i=1}^{n} \frac{1}{n}(T(i-1)+T(n-i))
\end{aligned}
$$

## Analysis via Recurrence

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We derived:

$$
E[Q(A)] \leq n+\sum_{i=1}^{n} \frac{1}{n}(T(i-1)+T(n-i))
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Note that above holds for any $\boldsymbol{A}$ of size $\boldsymbol{n}$. Therefore

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T(n)=\max _{A:|A|=n} \mathrm{E}[Q(A)] \leq n+\sum_{i=1}^{n} \frac{1}{n}(T(i-1)+T(n-i)) .
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## Solving the Recurrence

$$
T(n) \leq n+\sum_{i=1}^{n} \frac{1}{n}(T(i-1)+T(n-i))
$$

with base case $T(1)=0$.

## Lemma

$$
T(n)=O(n \log n) .
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with base case $T(1)=0$.
Lemma

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T(n)=O(n \log n) .
$$

## Proof.

(Guess and) Verify by induction.

## Part II

## Slick analysis of QuickSort

## A Slick Analysis of QuickSort

$Q(A)$ : number of comparisons done on input array $\boldsymbol{A}$
(1) Rank of an element is its position in the sorted $\boldsymbol{A}$.
(2) $R_{i j}$ : event that rank $i$ element is compared with rank $j$ element, for $\mathbf{1} \leq i<j<\boldsymbol{n}$.

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(3) $X_{i j}$ : the indicator random variable for $R_{i j}$. That is, $X_{i j}=1$ if rank $\boldsymbol{i}$ is compared with rank $\boldsymbol{j}$ element, otherwise $\mathbf{0}$.

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$$
Q(A)=\sum_{1 \leq i<j \leq n} X_{i j}
$$

and hence by linearity of expectation,

$$
\mathrm{E}[Q(A)]=\sum_{1 \leq i<j \leq n} \mathrm{E}\left[X_{i j}\right]=\sum_{1 \leq i<j \leq n} \operatorname{Pr}\left[R_{i j}\right] .
$$

## A Slick Analysis of QuickSort

$R_{i j}=$ rank $i$ element is compared with rank $j$ element.
Question: What is $\operatorname{Pr}\left[R_{i j}\right]$ ?

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With ranks: | 7 | 5 | 9 | 1 | 3 | 4 | 8 | 6 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 6 | 4 | 8 | 1 | 2 | 3 | 7 | 5 |

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As such, probability of comparing 5 to 8 is $\operatorname{Pr}\left[R_{4,7}\right]$.

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(1) If pivot too small (say 3 [rank 2]). Partition and call recursively: | 7 | 5 | 9 | 1 | 3 | 4 | 8 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |



Decision if to compare 5 to $\mathbf{8}$ is moved to subproblem.

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(1) If pivot too small (say 3 [rank 2]). Partition and call recursively:

| 7 | 5 | 9 | 1 | 3 | 4 | 8 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |



Decision if to compare 5 to $\mathbf{8}$ is moved to subproblem.
(2) If pivot too large (say 9 [rank 8]):


Decision if to compare 5 to $\mathbf{8}$ moved to subproblem.

## A Slick Analysis of QuickSort

## Question: What is $\operatorname{Pr}\left[\mathrm{R}_{\mathrm{i}, \mathrm{j}}\right]$ ?

| 7 | 5 | 9 | 1 | 3 | 4 | 8 | 6 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 6 | 4 | 8 | 1 | 2 | 3 | 7 | 5 |

As such, probability of comparing $\mathbf{5}$ to $\mathbf{8}$ is $\operatorname{Pr}\left[R_{4,7}\right]$.
(1) If pivot is $\mathbf{5}$ (rank 4). Bingo!

$$
\begin{array}{|l|l|l|l|l|l|l}
\hline 7 & \mid & 9 & 1 & 3 & 4 & 8 \\
\hline
\end{array}
$$



## A Slick Analysis of QuickSort

Question: What is $\operatorname{Pr}\left[\mathrm{R}_{\mathrm{i}, \mathrm{j}}\right]$ ?

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As such, probability of comparing 5 to 8 is $\operatorname{Pr}\left[R_{4,7}\right]$.
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$$
\begin{array}{|l|l|l|l|l|l|l|}
\hline 7 & 9 & 1 & 3 & 4 & 8 & 6 \\
\hline
\end{array}
$$

$$
\Longrightarrow \begin{array}{|l|l|l|l|l|l|l|l|}
\hline 1 & 3 & 4 & 5 & 7 & 9 & 8 & 6 \\
\hline
\end{array}
$$

(2) If pivot is $\mathbf{8}$ (rank 7). Bingo!

$$
\begin{array}{|l|l|l|l|l|l|l|}
\hline 7 & 5 & 9 & 1 & 3 & 4 & \begin{array}{|l|l|l|l|l|l|l|}
\hline 6 \\
\hline
\end{array} \Longrightarrow \begin{array}{|l|l|l|l|l|l|}
\hline 7 & 5 & 1 & 3 & 4 & 6 \\
\hline
\end{array} \\
\hline
\end{array}
$$

## A Slick Analysis of QuickSort

## Question: What is $\operatorname{Pr}\left[\mathrm{R}_{\mathrm{i}, \mathrm{j}}\right]$ ?

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As such, probability of comparing 5 to $\mathbf{8}$ is $\operatorname{Pr}\left[R_{4,7}\right]$.
(1) If pivot is 5 (rank 4). Bingo!

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\end{array} \Longrightarrow \begin{array}{|l|l|l|l|l|l|l|l|l|}
\hline 1 & 3 & 4 & 5 & 7 & 9 & 8 & 6 \\
\hline
\end{array}
$$

(2) If pivot is 8 (rank 7). Bingo!

(3) If pivot in between the two numbers (say 6 [rank 5]):

| 7 | 5 | 9 | 1 | 3 | 4 | 8 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |

$$
\Longrightarrow \begin{array}{|l|l|l|l|l|l|l|l|}
\hline 5 & 1 & 3 & 4 & 6 & 7 & 8 & 9 \\
\hline
\end{array}
$$

5 and 8 will never be compared to each other.

## A Slick Analysis of QuickSort

## Question: What is $\operatorname{Pr}\left[\mathrm{R}_{\mathrm{i}, \mathrm{j}}\right]$ ?

## Conclusion:

$\boldsymbol{R}_{\boldsymbol{i}, \boldsymbol{j}}$ happens if and only if: $i$ th or $j$ th ranked element is the first pivot out of $i$ th to $j$ th ranked elements.
$\operatorname{Pr}\left[R_{i, j}\right]=\operatorname{Pr}[i$ th or $j$ th ranked element is the pivot pivot has rank in $\{i, i+\mathbf{1} \ldots, \boldsymbol{j}-\mathbf{1}, \boldsymbol{j}\}$ ]

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There are $k=j-i+1$ relevant elements.

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There are $k=j-i+1$ relevant elements.

$$
\operatorname{Pr}\left[R_{i, j}\right]=\frac{2}{k}=\frac{2}{j-i+1}
$$

## A Slick Analysis of QuickSort

Question: What is $\operatorname{Pr}\left[R_{i j}\right]$ ?
Lemma
$\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1}$.

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## Proof.

Let $a_{1}, \ldots, a_{i}, \ldots, a_{j}, \ldots, a_{\boldsymbol{n}}$ be elements of $\boldsymbol{A}$ in sorted order.
Let $S=\left\{a_{i}, a_{i+1}, \ldots, a_{j}\right\}$

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Observation: If pivot is chosen outside $S$ then all of $S$ either in left array or right array.

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Observation: $\boldsymbol{a}_{\boldsymbol{i}}$ and $\boldsymbol{a}_{\boldsymbol{j}}$ separated when a pivot is chosen from $\boldsymbol{S}$ for the first time. Once separated no comparison.

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Observation: $\boldsymbol{a}_{\boldsymbol{i}}$ is compared with $\boldsymbol{a}_{\boldsymbol{j}}$ if and only if either $\boldsymbol{a}_{\boldsymbol{i}}$ or $\boldsymbol{a}_{\boldsymbol{j}}$ is chosen as a pivot from $S$ at separation...

## A Slick Analysis of QuickSort

## Continued...

Lemma

$$
\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1} .
$$

## Proof.

Let $a_{1}, \ldots, a_{i}, \ldots, a_{j}, \ldots, a_{n}$ be sort of $\boldsymbol{A}$. Let
$S=\left\{a_{i}, a_{i+1}, \ldots, a_{j}\right\}$
Observation: $\boldsymbol{a}_{\boldsymbol{i}}$ is compared with $\boldsymbol{a}_{\boldsymbol{j}}$ if and only if either $\boldsymbol{a}_{\boldsymbol{i}}$ or $\boldsymbol{a}_{\boldsymbol{j}}$ is chosen as a pivot from $S$ at separation.
Observation: Given that pivot is chosen from $S$ the probability that it is $a_{i}$ or $a_{j}$ is exactly $2 /|S|=2 /(j-i+1)$ since the pivot is chosen uniformly at random from the array.

## A Slick Analysis of QuickSort

Continued...

$$
\mathrm{E}[Q(A)]=\sum_{1 \leq i<j \leq n} \mathrm{E}\left[X_{i j}\right]=\sum_{1 \leq i<j \leq n} \operatorname{Pr}\left[R_{i j}\right] .
$$

## Lemma

$$
\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1} .
$$

## A Slick Analysis of QuickSort

 Continued...
## Lemma <br> $\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1}$.

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Continued...

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$$

## A Slick Analysis of QuickSort

Continued...

## Lemma

$$
\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1} .
$$

$$
\begin{aligned}
\mathrm{E}[Q(A)] & =\sum_{1 \leq i<j \leq n} \frac{2}{j-i+1} \\
& =\sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{2}{j-i+1}
\end{aligned}
$$

## A Slick Analysis of QuickSort

Continued...

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## A Slick Analysis of QuickSort

Continued...

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$$

$$
E[Q(A)]=2 \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{1}{j-i+1}
$$

## A Slick Analysis of QuickSort

Continued...

## Lemma

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\operatorname{Pr}\left[R_{i j}\right]=\frac{2}{j-i+1} .
$$

$$
\mathrm{E}[Q(A)]=2 \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{1}{j-i+1}=2 \sum_{i=1}^{n-1} \sum_{\Delta=2}^{n-i+1} \frac{1}{\Delta}
$$

## A Slick Analysis of QuickSort

## Continued...

## Lemma

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$$

$$
\begin{aligned}
E[Q(A)] & =2 \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{1}{j-i+1}=2 \sum_{i=1}^{n-1} \sum_{\Delta=2}^{n-i+1} \frac{1}{\Delta} \\
& =2 \sum_{i=1}^{n-1}\left(H_{n-i+1}-1\right) \leq 2 \sum_{1 \leq i<n} H_{n}
\end{aligned}
$$

$$
H_{k}=\sum_{i=1}^{k} \frac{1}{i}=\Theta(\log k)
$$

## A Slick Analysis of QuickSort

## Continued...

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& =2 \sum_{i=1}^{n-1}\left(H_{n-i+1}-1\right) \leq 2 \sum_{1 \leq i<n} H_{n} \\
& \leq 2 n H_{n}=O(n \log n)
\end{aligned}
$$

$$
H_{k}=\sum_{i=1}^{k} \frac{1}{i}=\Theta(\log k)
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## Part III

## Inequalities

## Massive randomness.. Is not that random.

Consider flipping a fair coin $\boldsymbol{n}$ times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: $k$ w.p. $\binom{n}{k}^{1 / 2 n}$.

$\mathrm{n}=2$

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## Massive randomness.. Is not that random.



This is known as concentration of mass.
This is a very special case of the law of large numbers.

## Side note...

Law of large numbers (weakest form)...

## Informal statement of law of large numbers

For $\boldsymbol{n}$ large enough, the middle portion of the binomial distribution looks like (converges to) the normal/Gaussian distribution.


## Massive randomness.. Is not that random.

## Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

## Massive randomness.. Is not that random.

## Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

Use of well known inequalities in analysis.

## Randomized QuickSort: A possible analysis

## Analysis

- Random variable $Q=\#$ comparisons made by randomized QuickSort on an array of $\boldsymbol{n}$ elements.


## Randomized QuickSort: A possible analysis

## Analysis

- Random variable $Q=\#$ comparisons made by randomized QuickSort on an array of $\boldsymbol{n}$ elements.
- Suppose $\operatorname{Pr}[Q \geq 10 n l g n] \leq c$. Also we know that $Q \leq n^{2}$.


## Randomized QuickSort: A possible analysis

## Analysis

- Random variable $Q=$ \#comparisons made by randomized

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## Question:

How to find $c$, or in other words bound $\operatorname{Pr}[Q \geq 10 n \log n]$ ?

## Markov's Inequality

## Markov's inequality

Let $\boldsymbol{X}$ be a non-negative random variable over a probability space $(\Omega, \operatorname{Pr})$. For any $\boldsymbol{a}>0$,

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\operatorname{Pr}[X \geq a] \leq \frac{E[X]}{a}
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Proof:

$$
\begin{aligned}
\mathrm{E}[X] & =\sum_{\omega \in \Omega} X(\omega) \operatorname{Pr}[\omega] \\
& \geq \sum_{\omega \in \Omega, X(\omega) \geq a} X(\omega) \operatorname{Pr}[\omega] \\
& \geq a \sum_{\omega \in \Omega, x(\omega) \geq a} \operatorname{Pr}[\omega] \\
& =a \operatorname{Pr}[X \geq a]
\end{aligned}
$$

Markov's Inequality: Proof by Picture


## Example: Balls in a bin

- $n$ black and white balls in a bin.
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## Question

How large $\boldsymbol{k}$ needs to be before our estimated value $\boldsymbol{p}$ is close to $\boldsymbol{p}^{*}$ ?

## Example: Balls in a bin

A rough estimate through Markov's inequality.

## Lemma

For any $k \geq 1$ and $p=B / k, \operatorname{Pr}\left[p \geq 2 p^{*}\right] \leq \frac{1}{2}$

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- For each $\mathbf{1} \leq \boldsymbol{i} \leq \boldsymbol{k}$ define random variable $\boldsymbol{X}_{\boldsymbol{i}}$, which is $\mathbf{1}$ if $\boldsymbol{i}^{\text {th }}$ ball is black, otherwise $\mathbf{0}$.
- $\mathrm{E}\left[X_{i}\right]=\operatorname{Pr}\left[X_{i}=1\right]=p^{*}$.


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- $B=\sum_{i=1}^{k} X_{i}$, then $\mathrm{E}[B]=\sum_{i=1}^{k} \mathrm{E}\left[X_{i}\right]=k p^{*}$.
- Markov's inequality gives, $\operatorname{Pr}\left[p \geq 2 p^{*}\right]=$

$$
\operatorname{Pr}\left[\frac{B}{k} \geq 2 p^{*}\right]=\operatorname{Pr}\left[B \geq 2 k p^{*}\right]=\operatorname{Pr}[B \geq 2 \mathrm{E}[B]] \leq \frac{1}{2}
$$

## Chebyshev's Inequality: Variance

## Variance

Given a random variable $\boldsymbol{X}$ over probability space ( $\Omega, \mathrm{Pr}$ ), variance of $\boldsymbol{X}$ is the measure of how much does it deviate from its mean value. Formally, $\operatorname{Var}(X)=\mathbf{E}\left[(X-E[X])^{2}\right]=\mathbf{E}\left[X^{2}\right]-E[X]^{2}$

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## Intuitive Derivation

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\operatorname{Var}(X) & =\mathrm{E}[Y] \\
& =\mathrm{E}\left[X^{2}\right]-2 \mathrm{E}[X] \mathrm{E}[X]+\mathrm{E}[X]^{2} \\
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## Chebyshev's Inequality: Variance

## Independence

Random variables $X$ and $Y$ are called mutually independent if

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\forall x, y \in \mathbb{R}, \operatorname{Pr}[X=x \wedge Y=y]=\operatorname{Pr}[X=x] \operatorname{Pr}[Y=y]
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If $X$ and $Y$ are independent random variables then $\operatorname{Var}(X+Y)=\operatorname{Var}(X)+\operatorname{Var}(Y)$.

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$Y=(X-E[X])^{2}$ is a non-negative random variable. Apply Markov's Inequality to $Y$ for $a^{2}$.

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\begin{aligned}
\left.\operatorname{Pr}\left[Y \geq a^{2}\right] \leq \mathrm{E} Y\right] / a^{2} & \Leftrightarrow \operatorname{Pr}\left[(X-\mathrm{E}[X])^{2} \geq a^{2}\right] \leq \operatorname{Var}(X) / a^{2} \\
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\end{aligned}
$$

$$
\begin{aligned}
& \operatorname{Pr}[X \leq \mathrm{E}[X]-a] \leq \operatorname{Var}(X) / a^{2} \text { AND } \\
& \operatorname{Pr}[X \geq \mathrm{E}[X]+a] \leq \operatorname{Var}(X) / a^{2}
\end{aligned}
$$

## Example:Balls in a bin (contd)

## Lemma

For $\mathbf{0}<\epsilon<\mathbf{1}, k \geq 1$ and $p=B / k, \operatorname{Pr}\left[\left|p-p^{*}\right| \geq \epsilon\right] \leq 1 / k \epsilon^{2}$.

## Proof.

- Recall: $X_{i}$ is $\mathbf{1}$ if $\boldsymbol{i}^{\text {th }}$ ball is black, else $0, B=\sum_{i=1}^{k} X_{i}$. $\mathrm{E}\left[X_{i}\right]=p^{*}, \mathrm{E}[B]=k p^{*} . p=B / k$.


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\begin{aligned}
\operatorname{Pr}\left[\left|p-p^{*}\right| \geq \epsilon\right] & =\operatorname{Pr}\left[\left|B / k-p^{*}\right| \geq \epsilon\right] \\
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(\text { Chebyshev }) & \leq \operatorname{Var}(B) / k^{2} \epsilon^{2}=k p^{*}\left(1-p^{*}\right) / k^{2} \epsilon^{2} \\
& <1 / k \epsilon^{2}
\end{aligned}
$$

## Chernoff Bound

## Lemma

Let $X_{1}, \ldots, X_{k}$ be $\boldsymbol{k}$ independent binary random variables such that, for each $i \in[1, k], X_{i}$ equals 1 w.p. $p_{i}$, and 0 w.p. $\left(1-p_{i}\right)$.

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For any $\mathbf{0}<\boldsymbol{\delta}<\mathbf{1}$, it holds that:

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\operatorname{Pr}[|X-\mu| \geq \delta \mu] \leq 2 e^{\frac{-\delta^{2} \mu}{3}}
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$$
\begin{gathered}
\operatorname{Pr}[|X-\mu| \geq \delta \mu] \leq 2 e^{\frac{-\delta^{2} \mu}{3}} \\
\operatorname{Pr}[X \geq(1+\delta) \mu] \leq e^{\frac{-\delta^{2} \mu}{3}} \text { and } \operatorname{Pr}[X \leq(1-\delta) \mu] \leq e^{\frac{-\delta^{2} \mu}{2}}
\end{gathered}
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Let $X_{1}, \ldots, X_{k}$ be $\boldsymbol{k}$ independent binary random variables such that, for each $i \in[1, k], X_{i}$ equals 1 w.p. $p_{i}$, and 0 w.p. $\left(1-p_{i}\right)$. Let $X=\sum_{i=1}^{k} X_{i}$ and $\mu=\mathrm{E}[X]=\sum_{i} p_{i}$.

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## Proof.

In notes!

## Example:Balls in a bin (Contd.)

## Lemma

For any $\mathbf{0}<\epsilon<\mathbf{1}$, and $k \geq \mathbf{1}, \operatorname{Pr}\left[\left|p-p^{*}\right| \geq \epsilon\right] \leq 2 e^{-\frac{k \epsilon^{2}}{3}}$.

## Proof.

Recall: $\boldsymbol{X}_{\boldsymbol{i}}$ is $\mathbf{1}$ is $\boldsymbol{i}^{\text {th }}$ ball is black, else $\mathbf{0}$. $B=\sum_{i=1}^{k} X_{i} . \mathrm{E}\left[X_{i}\right]=p^{*}, \mathrm{E}[B]=k p^{*} . p=B / k$.

## Example:Balls in a bin (Contd.)

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$$
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& =\operatorname{Pr}\left[\left|B-k p^{*}\right| \geq k \epsilon\right] \\
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$$

(Chernoff) $\leq 2 e^{-\frac{\epsilon^{2}}{3 p^{*}} k p^{*}}=2 e^{-\frac{k \epsilon^{2}}{3 p^{*}}}$

$$
\left(p^{*} \leq 1\right) \leq 2 e^{-\frac{k c^{2}}{3}}
$$

## Example Summary

The problem was to estimate the fraction of black balls $p^{*}$ in a bin filled with white and black balls. Our estimate was $p=\frac{B}{k}$ instead, where out of $k$ draws (with replacement) $B$ balls turns out black.

## Markov's Inequality

For any $k \geq 1, \operatorname{Pr}\left[p \geq 2 p^{*}\right] \leq \frac{1}{2}$
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For any $0<\epsilon<1$, and $k \geq 1, \operatorname{Pr}\left[\left|p-p^{*}\right| \geq \epsilon\right] \leq 1 / k \epsilon^{2}$.

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## Part IV

## Randomized QuickSort (Contd.)

## Randomized QuickSort: Recall

Input: Array $\boldsymbol{A}$ of $\boldsymbol{n}$ numbers. Output: Numbers in sorted order.

## Randomized QuickSort

(1) Pick a pivot element uniformly at random from $\boldsymbol{A}$.
(2) Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
(3) Recursively sort the subarrays, and concatenate them.

Note: On every input randomized QuickSort takes $O(n \log n)$ time in expectation. On every input it may take $\boldsymbol{\Omega}\left(\boldsymbol{n}^{2}\right)$ time with some small probability.

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Question: With what probability it takes $O(n \log n)$ time?

## Randomized QuickSort: High Probability Analysis

## Informal Statement

Random variable $Q(A)=\#$ comparisons done by the algorithm. We will show that $\operatorname{Pr}[Q(A) \leq 32 n \ln n] \geq 1-1 / n^{3}$.

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Random variable $Q(A)=\#$ comparisons done by the algorithm. We will show that $\operatorname{Pr}[Q(A) \leq 32 n \ln n] \geq 1-1 / n^{3}$.

$$
\text { If } n=100 \text { then this gives } \operatorname{Pr}[Q(A) \leq 32 n \ln n] \geq 0.99999 \text {. }
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Outline of the proof

- $\boldsymbol{k}$ : depth of the recursion. Then $Q(A) \leq k n$.


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(1) Focus on a single element. Prove that it "participates" in
$>32 \ln \boldsymbol{n}$ levels with probability at most $1 / \boldsymbol{n}^{4}$.
(2) By union bound, any of the $\boldsymbol{n}$ elements participates in $>32 \ln \boldsymbol{n}$ levels with probability at most


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(2) By union bound, any of the $\boldsymbol{n}$ elements participates in
$>32 \ln n$ levels with probability at most $1 / n^{3}$.
(3) Therefore, all elements participate in $\leq 32 \ln n$ w.p. $\left(1-1 / n^{3}\right)$.


## Randomized QuickSort: High Probability Analysis

- If $\boldsymbol{k}$ levels of recursion then $\boldsymbol{k} \boldsymbol{n}$ comparisons.


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- If $\boldsymbol{k}$ levels of recursion then $\boldsymbol{k} \boldsymbol{n}$ comparisons.
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- $S_{1}=A$ and $S_{k}=\{s\}$.


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- $S_{1}=A$ and $S_{k}=\{s\}$.
- We call $s$ lucky in $i^{\text {th }}$ iteration, if balanced split:

$$
\left|S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right| \text { and }\left|S_{i} \backslash S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right| .
$$

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- Let $S_{i}$ be the partition containing $s$ at $i^{\text {th }}$ level.
- $S_{1}=A$ and $S_{k}=\{s\}$.
- We call $s$ lucky in $i^{\text {th }}$ iteration, if balanced split:

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\left|S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right| \text { and }\left|S_{i} \backslash S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right| .
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## Randomized QuickSort: High Probability Analysis

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- If $\rho=$ \#lucky rounds in first $k$ rounds, then
$\left|S_{k}\right| \leq(3 / 4)^{\rho} n$.
- For $\left|S_{k}\right|=1, \rho=\log _{4 / 3} n \leq 4 \ln n$ suffices.


## How many rounds before $4 \ln n$ lucky rounds?

$s$ lucky in round $i$ if $\left|S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right|$ and $\left|S_{i} \backslash S_{i+1}\right| \leq(3 / 4)\left|S_{i}\right|$

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(Chernoff) $\leq 2 e^{\frac{-\delta^{2} \mu}{2}}=2 e^{-\frac{9 k}{64}}$

$$
=2 e^{-4.5 \ln n} \leq \frac{1}{n^{4}}
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Q: How to increase the probability?

