# CS 473: Algorithms

#### Ruta Mehta

University of Illinois, Urbana-Champaign

Spring 2021

## CS 473: Algorithms, Spring 2021

# Inequalities & Randomized QuickSort

Lecture 8 Feb 18, 2021

Most slides are courtesy Prof. Chekuri

Ruta (UIUC)

CS473

### Outline

#### Slick Analysis of Randomized QuickSort

#### Concentration of Mass Around Mean

Markov's Inequality

Chebyshev's Inequality

Chernoff Bound

Randomized **QuickSort**: High Probability Analysis

# Part I

# Analysis of QuickSort

### Recall: Randomized QuickSort

#### Randomized QuickSort

- **1** Pick a pivot element *uniformly at random* from the array.
- Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- In the subarrays, and concatenate them.

#### Theorem

Expected running time of Randomized QuickSort on an array of size n is  $O(n \log n)$ .

- A: Given array with *n* distinct numbers.
- Q(A): number of comparisons of randomized QuickSort on A.
   Note that Q(A) is a random variable.
- **3**  $X_i$ : Random variable indicating if picked pivot has rank i in A.

 $A_{left}^{i}$  and  $A_{right}^{i}$  be the corresponding left and right subarrays.

$$Q(A) = n + \sum_{i=1}^{n} X_i \cdot \left(Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i)\right).$$

Exactly one non-zero  $X_i$ .  $E[X_i] = Pr[pivot has rank i] = 1/n$ .

## Independence of Random Variables

#### Lemma

Random variables  $X_i$  is independent of random variables  $Q(A_{left}^i)$  as well as  $Q(A_{right}^i)$ , i.e.

$$\mathbf{E} \begin{bmatrix} X_i \cdot Q(A_{left}^i) \end{bmatrix} = \mathbf{E} \begin{bmatrix} X_i \end{bmatrix} \mathbf{E} \begin{bmatrix} Q(A_{left}^i) \end{bmatrix}$$
$$\mathbf{E} \begin{bmatrix} X_i \cdot Q(A_{right}^i) \end{bmatrix} = \mathbf{E} \begin{bmatrix} X_i \end{bmatrix} \mathbf{E} \begin{bmatrix} Q(A_{right}^i) \end{bmatrix}$$

#### Proof.

This is because the algorithm, while recursing on  $Q(A_{left}^{i})$  and  $Q(A_{right}^{i})$  uses new random coin tosses that are independent of the coin tosses used to decide the first pivot. Only the latter decides value of  $X_{i}$ .

Ruta (UIUC)

 $T(n) = \max_{A:|A|=n} E[Q(A)]$  be the worst-case expected running time on arrays of size n.

We have, for any **A**:

$$Q(A) = n + \sum_{i=1}^{n} X_i \left( Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i) \right)$$

 $T(n) = \max_{A:|A|=n} E[Q(A)]$  be the worst-case expected running time on arrays of size n.

We have, for any **A**:

$$Q(A) = n + \sum_{i=1}^{n} X_i \left( Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i) \right)$$

By linearity of expectation, and independence random variables:

 $T(n) = \max_{A:|A|=n} E[Q(A)]$  be the worst-case expected running time on arrays of size n. We derived:

$$\mathsf{E}\Big[Q(A)\Big] \leq n + \sum_{i=1}^{n} \frac{1}{n} \left(T(i-1) + T(n-i)\right).$$

Note that above holds for any A of size n. Therefore

 $T(n) = \max_{A:|A|=n} \mathsf{E}[Q(A)] \leq$ 

 $T(n) = \max_{A:|A|=n} E[Q(A)]$  be the worst-case expected running time on arrays of size n. We derived:

$$\mathsf{E}\Big[Q(A)\Big] \leq n + \sum_{i=1}^{n} \frac{1}{n} \left(T(i-1) + T(n-i)\right).$$

Note that above holds for any A of size n. Therefore

$$T(n) = \max_{A:|A|=n} \mathbb{E}[Q(A)] \le n + \sum_{i=1}^{n} \frac{1}{n} (T(i-1) + T(n-i)).$$

## Solving the Recurrence

$$T(n) \leq n + \sum_{i=1}^{n} \frac{1}{n} (T(i-1) + T(n-i))$$

with base case T(1) = 0.

#### Lemma

 $T(n) = O(n \log n).$ 

# Solving the Recurrence

$$T(n) \leq n + \sum_{i=1}^{n} \frac{1}{n} (T(i-1) + T(n-i))$$

with base case T(1) = 0.

#### Lemma

 $T(n) = O(n \log n).$ 

#### Proof.

(Guess and) Verify by induction.

# Part II

# Slick analysis of QuickSort

Q(A): number of comparisons done on input array A

- Sank of an element is its position in the sorted A.
- R<sub>ij</sub>: event that rank *i* element is compared with rank *j* element, for 1 ≤ *i* < *j* < *n*.

Q(A): number of comparisons done on input array A

- Sank of an element is its position in the sorted A.
- R<sub>ij</sub>: event that rank *i* element is compared with rank *j* element, for 1 ≤ *i* < *j* < *n*.
- X<sub>ij</sub>: the indicator random variable for R<sub>ij</sub>. That is, X<sub>ij</sub> = 1 if rank *i* is compared with rank *j* element, otherwise 0.

Q(A) : number of comparisons done on input array A

- Sank of an element is its position in the sorted A.
- *R<sub>ij</sub>*: event that rank *i* element is compared with rank *j* element, for 1 ≤ *i* < *j* < *n*.
- X<sub>ij</sub>: the indicator random variable for R<sub>ij</sub>. That is, X<sub>ij</sub> = 1 if rank *i* is compared with rank *j* element, otherwise 0.

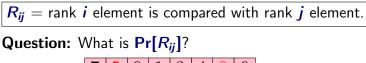
$$Q(A) = \sum_{1 \le i < j \le n} X_{ij}$$

and hence by linearity of expectation,

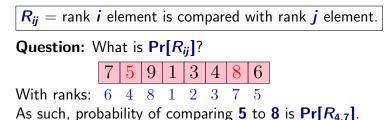
$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \le i < j \le n} \mathsf{E}\Big[X_{ij}\Big] = \sum_{1 \le i < j \le n} \mathsf{Pr}\Big[R_{ij}\Big].$$

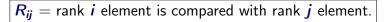
 $R_{ij}$  = rank *i* element is compared with rank *j* element.

Question: What is Pr[R<sub>ij</sub>]?



With ranks:  $6 \ 4 \ 8 \ 1 \ 2 \ 3 \ 7 \ 5$ 





**Question:** What is **Pr**[*R*<sub>ij</sub>]?

13486

With ranks:  $6 \ 4 \ 8 \ 1 \ 2 \ 3 \ 7 \ 5$ 

If pivot too small (say 3 [rank 2]). Partition and call recursively:

Decision if to compare **5** to **8** is moved to subproblem.

5 | 9



**Question:** What is **Pr**[*R*<sub>ij</sub>]?

With ranks:  $6 \ 4 \ 8 \ 1 \ 2 \ 3 \ 7 \ 5$ 

If pivot too small (say 3 [rank 2]). Partition and call recursively:

Decision if to compare 5 to 8 is moved to subproblem.

If pivot too large (say 9 [rank 8]):

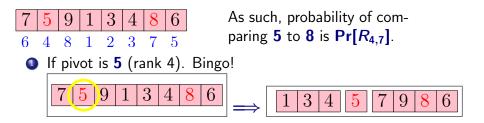
3 | 4

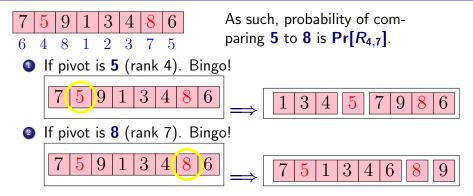
9 1 3 4 8 6

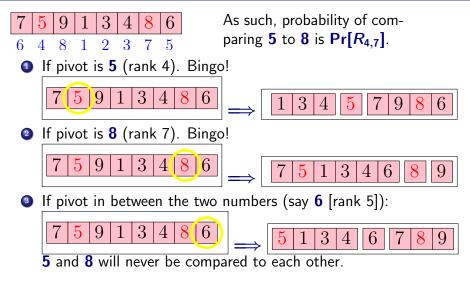
$$\overrightarrow{5} \implies \boxed{75134869}$$

Decision if to compare 5 to 8 moved to subproblem.

5







#### Conclusion:

**R**<sub>*i*,*j*</sub> happens if and only if:

*i*th or *j*th ranked element is the first pivot out of *i*th to *j*th ranked elements.

 $\Pr[R_{i,j}] = \Pr[i \text{th or } j \text{th ranked element is the pivot } | \\pivot has rank in \{i, i + 1, \dots, j - 1, j\}]$ 

#### Conclusion:

 $R_{i,j}$  happens if and only if:

*i*th or *j*th ranked element is the first pivot out of *i*th to *j*th ranked elements.

 $\Pr[R_{i,j}] = \Pr[i \text{th or } j \text{th ranked element is the pivot } |$ pivot has rank in  $\{i, i + 1, \dots, j - 1, j\}$ ]

There are k = j - i + 1 relevant elements.

#### Conclusion:

 $R_{i,j}$  happens if and only if:

*i*th or *j*th ranked element is the first pivot out of *i*th to *j*th ranked elements.

 $\Pr[R_{i,j}] = \Pr[i \text{th or } j \text{th ranked element is the pivot } |$ pivot has rank in  $\{i, i + 1, \dots, j - 1, j\}$ ]

There are k = j - i + 1 relevant elements.

$$\Pr\left[R_{i,j}\right] = \frac{2}{k} = \frac{2}{j-i+1}.$$

#### **Question:** What is **Pr**[*R*<sub>*ij*</sub>]?

#### Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

#### Question: What is **Pr**[*R*<sub>ij</sub>]?

# Lemma $\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$

#### Proof.

Let  $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$  be elements of A in sorted order. Let  $S = \{a_i, a_{i+1}, \ldots, a_j\}$ 

#### Question: What is **Pr**[*R*<sub>ij</sub>]?

#### Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

#### Proof.

Let  $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$  be elements of A in sorted order. Let  $S = \{a_i, a_{i+1}, \ldots, a_j\}$ **Observation:** If pivot is chosen outside S then all of S either in left array or right array.

#### Question: What is Pr[R<sub>ij</sub>]?

#### Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

#### Proof.

Let  $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$  be elements of A in sorted order. Let  $S = \{a_i, a_{i+1}, \ldots, a_j\}$ **Observation:** If pivot is chosen outside S then all of S either in left array or right array. **Observation:**  $a_i$  and  $a_j$  separated when a pivot is chosen from S for the first time. Once separated no comparison.

#### Question: What is Pr[R<sub>ij</sub>]?

#### Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

#### Proof.

Let  $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$  be elements of A in sorted order. Let  $S = \{a_i, a_{i+1}, \ldots, a_j\}$ **Observation:** If pivot is chosen outside S then all of S either in left array or right array. **Observation:**  $a_i$  and  $a_j$  separated when a pivot is chosen from S for the first time. Once separated no comparison. **Observation:**  $a_i$  is compared with  $a_j$  if and only if either  $a_i$  or  $a_j$  is chosen as a pivot from S at separation...

#### A Slick Analysis of **QuickSort** Continued...

#### Lemma

$$\Pr\left[R_{ij}\right] = \frac{2}{j-i+1}.$$

#### Proof.

Let  $a_1, \ldots, a_i, \ldots, a_j, \ldots, a_n$  be sort of A. Let  $S = \{a_i, a_{i+1}, \ldots, a_j\}$  **Observation:**  $a_i$  is compared with  $a_j$  if and only if either  $a_i$  or  $a_j$  is chosen as a pivot from S at separation. **Observation:** Given that pivot is chosen from S the probability that it is  $a_i$  or  $a_j$  is exactly 2/|S| = 2/(j - i + 1) since the pivot is chosen uniformly at random from the array.

#### A Slick Analysis of **QuickSort** Continued...

$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \leq i < j \leq n} \mathsf{E}[X_{ij}] = \sum_{1 \leq i < j \leq n} \mathsf{Pr}[R_{ij}].$$

Lemma	
$\Pr[R_{ij}] = \frac{2}{j-i+1}.$	

#### A Slick Analysis of **QuickSort** Continued...

#### Lemma

$$\Pr[R_{ij}] = \frac{2}{j-i+1}.$$

$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \le i < j \le n} \mathsf{Pr}\Big[R_{ij}\Big] = \sum_{1 \le i < j \le n} \frac{2}{j-i+1}$$

#### Lemma

$$\Pr[R_{ij}] = \frac{2}{j-i+1}.$$

$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \le i < j \le n} \frac{2}{j-i+1}$$

#### Lemma

$$\mathsf{E}\Big[Q(A)\Big] = \sum_{1 \le i < j \le n} \frac{2}{j-i+1}$$
$$= \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{2}{j-i+1}$$

#### Lemma

$$E[Q(A)] = \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \frac{2}{j-i+1}$$

#### Lemma

$$E[Q(A)] = 2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n}\frac{1}{j-i+1}$$

#### Lemma

$$\mathsf{E}\Big[Q(A)\Big] = 2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n}\frac{1}{j-i+1} = 2\sum_{i=1}^{n-1}\sum_{\Delta=2}^{n-i+1}\frac{1}{\Delta}$$

#### Lemma

$$\Pr[R_{ij}] = \frac{2}{j-i+1}$$

$$\mathsf{E}\Big[Q(A)\Big] = 2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n}\frac{1}{j-i+1} = 2\sum_{i=1}^{n-1}\sum_{\Delta=2}^{n-i+1}\frac{1}{\Delta}$$
$$= 2\sum_{i=1}^{n-1}(H_{n-i+1}-1) \leq 2\sum_{1\leq i< n}H_{n}$$

$$H_k = \sum_{i=1}^k \frac{1}{i} = \Theta(\log k)$$

#### Lemma

$$\Pr[R_{ij}] = \frac{2}{j-i+1}.$$

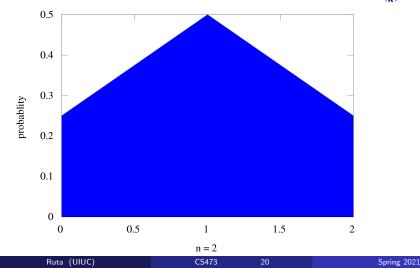
$$E\left[Q(A)\right] = 2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n}\frac{1}{j-i+1} = 2\sum_{i=1}^{n-1}\sum_{\Delta=2}^{n-i+1}\frac{1}{\Delta}$$
$$= 2\sum_{i=1}^{n-1}(H_{n-i+1}-1) \leq 2\sum_{1\leq i< n}H_{n}$$
$$\leq 2nH_{n} = O(n\log n)$$

### $H_k = \sum_{i=1}^k \frac{1}{i} = \Theta(\log k)$

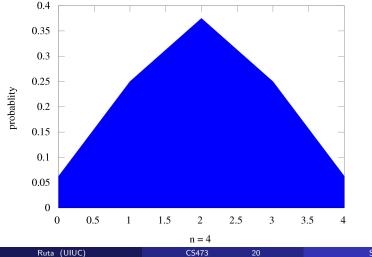
# Part III

# Inequalities

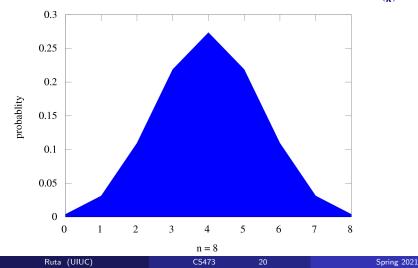
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



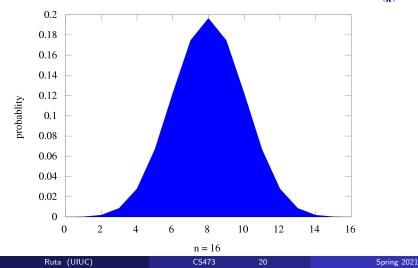
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



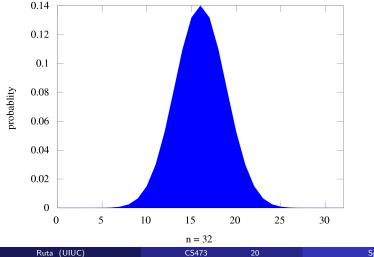
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



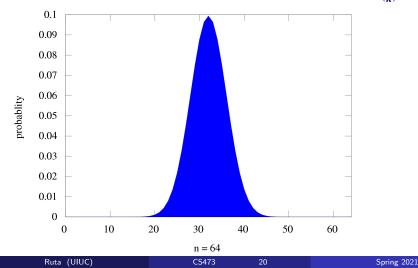
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



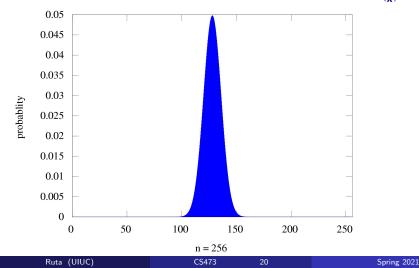
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



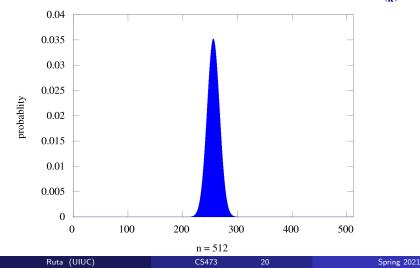
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



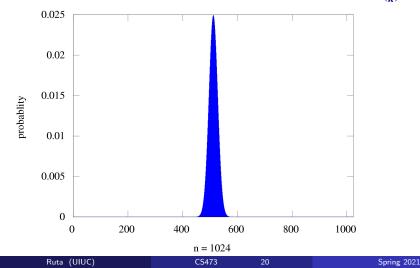
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



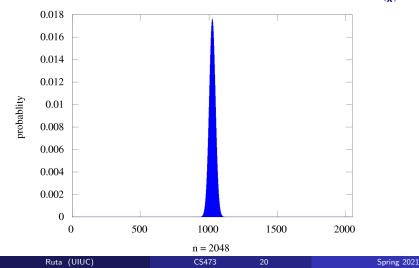
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



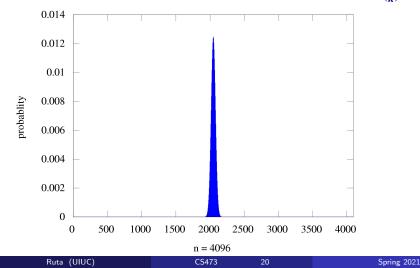
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



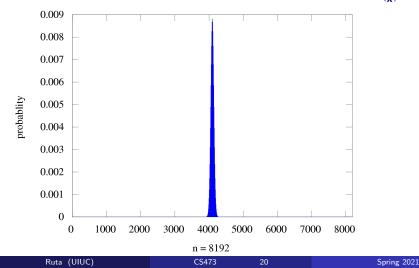
Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .

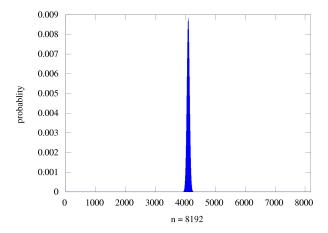


Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .



Consider flipping a fair coin *n* times independently, head gives 1, tail gives zero. How many 1s? Binomial distribution: k w.p.  $\binom{n}{k} \frac{1}{2^n}$ .

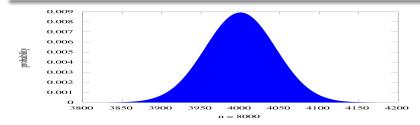




This is known as **concentration of mass**. This is a very special case of the **law of large numbers**.

#### Informal statement of law of large numbers

For n large enough, the middle portion of the binomial distribution looks like (converges to) the normal/Gaussian distribution.



#### Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

#### Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

Use of well known inequalities in analysis.

#### Analysis

 Random variable Q = #comparisons made by randomized QuickSort on an array of n elements.

#### Analysis

- Random variable Q = #comparisons made by randomized QuickSort on an array of n elements.
- Suppose  $\Pr[Q \ge 10 n lgn] \le c$ . Also we know that  $Q \le n^2$ .

#### Analysis

- Random variable Q = #comparisons made by randomized QuickSort on an array of n elements.
- Suppose  $\Pr[Q \ge 10 n lgn] \le c$ . Also we know that  $Q \le n^2$ .
- $E[Q] \le (10n \log n)(1-c) + n^2 c$

#### Analysis

- Random variable Q = #comparisons made by randomized QuickSort on an array of n elements.
- Suppose  $\Pr[Q \ge 10 n lgn] \le c$ . Also we know that  $Q \le n^2$ .
- $E[Q] \le (10n \log n)(1-c) + n^2 c$

### Question:

How to find c, or in other words bound  $\Pr[Q \ge 10n \log n]$ ?

# Markov's Inequality

#### Markov's inequality

Let X be a **non-negative** random variable over a probability space  $(\Omega, \Pr)$ . For any a > 0,

$$\Pr[X \ge a] \le \frac{\mathsf{E}[X]}{a}$$

# Markov's Inequality

#### Markov's inequality

Let X be a **non-negative** random variable over a probability space  $(\Omega, \Pr)$ . For any a > 0,

$$\Pr[X \ge a] \le \frac{\mathsf{E}[X]}{a}$$

### Proof:

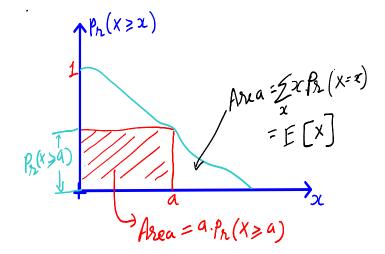
$$E[X] = \sum_{\omega \in \Omega} X(\omega) \Pr[\omega]$$
  

$$\geq \sum_{\omega \in \Omega, \ X(\omega) \geq a} X(\omega) \Pr[\omega]$$
  

$$\geq a \sum_{\omega \in \Omega, \ X(\omega) \geq a} \Pr[\omega]$$
  

$$= a \Pr[X \geq a]$$

# Markov's Inequality: Proof by Picture



- *n* black and white balls in a bin.
- We wish to estimate the fraction of black balls. Lets say it is  $p^*$ .

- *n* black and white balls in a bin.
- We wish to estimate the fraction of black balls. Lets say it is *p*\*.
- An approach: Draw k balls with replacement. If B are black then output  $p = \frac{B}{k}$ .

- *n* black and white balls in a bin.
- We wish to estimate the fraction of black balls. Lets say it is *p*\*.
- An approach: Draw k balls with replacement. If B are black then output  $p = \frac{B}{k}$ .

#### Question

How large k needs to be before our estimated value p is close to  $p^*$ ?

A rough estimate through Markov's inequality.

#### Lemma

### For any $k \geq 1$ and p = B/k, $\Pr[p \geq 2p^*] \leq \frac{1}{2}$

A rough estimate through Markov's inequality.

#### Lemma

```
For any k \geq 1 and p = {}^B/k, \Pr[p \geq 2p^*] \leq \frac{1}{2}
```

#### Proof.

- For each 1 ≤ i ≤ k define random variable X<sub>i</sub>, which is 1 if i<sup>th</sup> ball is black, otherwise 0.
- $E[X_i] = Pr[X_i = 1] = p^*$ .

# Example: Balls in a bin

A rough estimate through Markov's inequality.

#### Lemma

```
For any k \geq 1 and p = {}^B/k, \Pr[p \geq 2p^*] \leq \frac{1}{2}
```

### Proof.

- For each 1 ≤ i ≤ k define random variable X<sub>i</sub>, which is 1 if i<sup>th</sup> ball is black, otherwise 0.
- $E[X_i] = Pr[X_i = 1] = p^*$ .
- $B = \sum_{i=1}^{k} X_i$ , then  $E[B] = \sum_{i=1}^{k} E[X_i] = kp^*$ .

# Example: Balls in a bin

A rough estimate through Markov's inequality.

#### Lemma

For any 
$$k\geq 1$$
 and  $p={}^{B}\!/{}_{k},$   $\Pr[p\geq 2p^{*}]\leq rac{1}{2}$ 

### Proof.

- For each 1 ≤ i ≤ k define random variable X<sub>i</sub>, which is 1 if i<sup>th</sup> ball is black, otherwise 0.
- $E[X_i] = Pr[X_i = 1] = p^*$ .
- $B = \sum_{i=1}^{k} X_i$ , then  $E[B] = \sum_{i=1}^{k} E[X_i] = kp^*$ .
- Markov's inequality gives,  $\Pr[p \ge 2p^*] =$

$$\Pr\left[\frac{B}{k} \ge 2p^*\right] = \Pr[B \ge 2kp^*] = \Pr[B \ge 2\operatorname{E}[B]] \le \frac{1}{2}$$

#### Variance

Given a random variable X over probability space  $(\Omega, Pr)$ , variance of X is the measure of how much does it deviate from its mean value. Formally,  $Var(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$ 

#### Variance

Given a random variable X over probability space  $(\Omega, Pr)$ , variance of X is the measure of how much does it deviate from its mean value. Formally,  $Var(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$ 

### Intuitive Derivation

Define  $Y = (X - E[X])^2 = X^2 - 2X E[X] + E[X]^2$ .

#### Variance

Given a random variable X over probability space  $(\Omega, Pr)$ , variance of X is the measure of how much does it deviate from its mean value. Formally,  $Var(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$ 

#### Intuitive Derivation

Define 
$$Y = (X - E[X])^2 = X^2 - 2X E[X] + E[X]^2$$
.

$$Var(X) = E[Y] = E[X^{2}] - 2 E[X] E[X] + E[X]^{2} = E[X^{2}] - E[X]^{2}$$

#### Independence

Random variables X and Y are called mutually independent if  $\forall x, y \in \mathbb{R}, \ \Pr[X = x \land Y = y] = \Pr[X = x] \Pr[Y = y]$ 

#### Lemma

If X and Y are independent random variables then Var(X + Y) = Var(X) + Var(Y).

#### Independence

Random variables X and Y are called mutually independent if  $\forall x, y \in \mathbb{R}, \ \Pr[X = x \land Y = y] = \Pr[X = x] \Pr[Y = y]$ 

#### Lemma

If X and Y are independent random variables then Var(X + Y) = Var(X) + Var(Y).

#### Lemma

If X and Y are mutually independent, then E[XY] = E[X] E[Y].

# Chebyshev's Inequality

#### Chebyshev's Inequality

Given  $a \ge 0$ ,  $\Pr[|X - E[X]| \ge a] \le \frac{Var(X)}{a^2}$ 

# Chebyshev's Inequality

### Chebyshev's Inequality

Given 
$$a \ge 0$$
,  $\Pr[|X - E[X]| \ge a] \le \frac{Var(X)}{a^2}$ 

#### Proof.

 $Y = (X - E[X])^2$  is a non-negative random variable. Apply Markov's Inequality to Y for  $a^2$ .

# $\begin{aligned} \mathsf{Pr}\big[Y \geq a^2\big] \leq \mathsf{E}^{[Y]}/a^2 & \Leftrightarrow \quad \mathsf{Pr}\big[(X - \mathsf{E}[X])^2 \geq a^2\big] \leq \frac{\mathsf{Var}(X)}{a^2} \\ & \Leftrightarrow \quad \mathsf{Pr}\big[|X - \mathsf{E}[X]| \geq a\big] \leq \frac{\mathsf{Var}(X)}{a^2} \end{aligned}$

# Chebyshev's Inequality

### Chebyshev's Inequality

Given 
$$a \ge 0$$
,  $\Pr[|X - E[X]| \ge a] \le \frac{Var(X)}{a^2}$ 

#### Proof.

 $Y = (X - E[X])^2$  is a non-negative random variable. Apply Markov's Inequality to Y for  $a^2$ .

 $\begin{aligned} \mathsf{Pr}\big[Y \geq a^2\big] \leq \mathsf{E}^{[Y]}/_{a^2} & \Leftrightarrow \quad \mathsf{Pr}\big[(X - \mathsf{E}[X])^2 \geq a^2\big] \leq \frac{\mathsf{Var}(X)}{a^2} \\ & \Leftrightarrow \quad \mathsf{Pr}\big[|X - \mathsf{E}[X]| \geq a\big] \leq \frac{\mathsf{Var}(X)}{a^2} \end{aligned}$ 

$$\begin{aligned} &\mathsf{Pr}[X \leq \mathsf{E}[X] - a] \leq \frac{\operatorname{Var}(X)}{a^2} \text{ AND} \\ &\mathsf{Pr}[X \geq \mathsf{E}[X] + a] \leq \frac{\operatorname{Var}(X)}{a^2} \end{aligned}$$

#### Lemma

For 
$$0 < \epsilon < 1$$
,  $k \ge 1$  and  $p = B/k$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Proof.

• Recall:  $X_i$  is 1 if  $i^{th}$  ball is black, else 0,  $B = \sum_{i=1}^k X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ .  $p = {}^B/k$ .

#### Lemma

For 
$$0 < \epsilon < 1$$
,  $k \ge 1$  and  $p = B/k$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Proof.

• Recall:  $X_i$  is 1 if  $i^{th}$  ball is black, else 0,  $B = \sum_{i=1}^k X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ .  $p = {}^B/k$ .

#### Lemma

For 
$$0 < \epsilon < 1$$
,  $k \ge 1$  and  $p = B/k$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Proof.

- Recall:  $X_i$  is 1 if  $i^{th}$  ball is black, else 0,  $B = \sum_{i=1}^{k} X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ . p = B/k.
- $Var(B) = \sum_{i} Var(X_i) = kp^*(1 p^*)$  (Exercise)

#### Lemma

For 
$$0 < \epsilon < 1$$
,  $k \ge 1$  and  $p = B/k$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Proof.

Recall: X<sub>i</sub> is 1 if i<sup>th</sup> ball is black, else 0, B = ∑<sub>i=1</sub><sup>k</sup> X<sub>i</sub>. E[X<sub>i</sub>] = p<sup>\*</sup>, E[B] = kp<sup>\*</sup>. p = <sup>B</sup>/k.
Var(B) = ∑<sub>i</sub> Var(X<sub>i</sub>) = kp<sup>\*</sup>(1 - p<sup>\*</sup>) (Exercise)

$$\Pr[|p - p^*| \ge \epsilon] = \Pr[|B/k - p^*| \ge \epsilon] \\ = \Pr[|B - kp^*| \ge k\epsilon]$$

#### Lemma

For 
$$0 < \epsilon < 1$$
,  $k \ge 1$  and  $p = B/k$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Proof.

 Recall: X<sub>i</sub> is 1 if i<sup>th</sup> ball is black, else 0, B = ∑<sub>i=1</sub><sup>k</sup> X<sub>i</sub>. E[X<sub>i</sub>] = p\*, E[B] = kp\*. p = <sup>B</sup>/k.
 Var(B) = ∑<sub>i</sub> Var(X<sub>i</sub>) = kp\*(1 - p\*) (Exercise)

$$\begin{aligned} \Pr[|p - p^*| \ge \epsilon] &= \Pr[|B/k - p^*| \ge \epsilon] \\ &= \Pr[|B - kp^*| \ge k\epsilon] \\ (\text{Chebyshev}) &\le \frac{\operatorname{Var}(B)}{k^2 \epsilon^2} = \frac{kp^*(1 - p^*)}{k^2 \epsilon^2} \\ &< \frac{1}{k\epsilon^2} \end{aligned}$$

#### Lemma

Let  $X_1, \ldots, X_k$  be k independent binary random variables such that, for each  $i \in [1, k]$ ,  $X_i$  equals 1 w.p.  $p_i$ , and 0 w.p.  $(1 - p_i)$ .

#### Lemma

Let  $X_1, \ldots, X_k$  be k independent binary random variables such that, for each  $i \in [1, k]$ ,  $X_i$  equals 1 w.p.  $p_i$ , and 0 w.p.  $(1 - p_i)$ . Let  $X = \sum_{i=1}^k X_i$  and  $\mu = \mathbf{E}[X] = \sum_i p_i$ .

#### Lemma

Let  $X_1, \ldots, X_k$  be k independent binary random variables such that, for each  $i \in [1, k]$ ,  $X_i$  equals 1 w.p.  $p_i$ , and 0 w.p.  $(1 - p_i)$ . Let  $X = \sum_{i=1}^k X_i$  and  $\mu = \mathbf{E}[X] = \sum_i p_i$ .

For any  $0 < \delta < 1$ , it holds that:

$$\Pr[|X - \mu| \ge \delta\mu] \le 2e^{rac{-\delta^2\mu}{3}}$$

#### Lemma

Let  $X_1, \ldots, X_k$  be k independent binary random variables such that, for each  $i \in [1, k]$ ,  $X_i$  equals 1 w.p.  $p_i$ , and 0 w.p.  $(1 - p_i)$ . Let  $X = \sum_{i=1}^k X_i$  and  $\mu = \mathbf{E}[X] = \sum_i p_i$ .

For any  $0 < \delta < 1$ , it holds that:

$$\mathsf{Pr}[|m{X}-\mu|\geq\delta\mu]\leq 2e^{rac{-\delta^{2}\mu}{3}}$$

 $\Pr[X \ge (1+\delta)\mu] \le e^{rac{-\delta^2\mu}{3}}$  and  $\Pr[X \le (1-\delta)\mu] \le e^{rac{-\delta^2\mu}{2}}$ 

#### Lemma

Let  $X_1, \ldots, X_k$  be k independent binary random variables such that, for each  $i \in [1, k]$ ,  $X_i$  equals 1 w.p.  $p_i$ , and 0 w.p.  $(1 - p_i)$ . Let  $X = \sum_{i=1}^k X_i$  and  $\mu = \mathbf{E}[X] = \sum_i p_i$ .

For any  $0 < \delta < 1$ , it holds that:

$$\mathsf{Pr}[|m{X}-\mu|\geq\delta\mu]\leq 2e^{rac{-\delta^{2}\mu}{3}}$$

 $\mathsf{Pr}[X \geq (1+\delta)\mu] \leq e^{rac{-\delta^2\mu}{3}}$  and  $\mathsf{Pr}[X \leq (1-\delta)\mu] \leq e^{rac{-\delta^2\mu}{2}}$ 

Proof.

In notes!

#### Lemma

For any 
$$0<\epsilon<1$$
, and  $k\geq 1$ ,  $\mathsf{Pr}[|p-p^*|\geq\epsilon]\leq 2e^{-rac{k\epsilon^2}{3}}$ 

### Proof.

Recall:  $X_i$  is 1 is  $i^{th}$  ball is black, else 0.  $B = \sum_{i=1}^{k} X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ . p = B/k. 2

#### Lemma

For any 
$$0<\epsilon<1$$
, and  $k\geq 1$ ,  $\mathsf{Pr}[|p-p^*|\geq\epsilon]\leq 2e^{-rac{k\epsilon^2}{3}}$ 

### Proof.

Recall:  $X_i$  is 1 is  $i^{th}$  ball is black, else 0.  $B = \sum_{i=1}^{k} X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ . p = B/k.  $Pr[|p - p^*| \ge \epsilon] = Pr[|\frac{B}{k} - p^*| \ge \epsilon]$   $= Pr[|B - kp^*| \ge k\epsilon]$  $= Pr[|B - kp^*| \ge (\frac{\epsilon}{p^*})kp^*]$ 

2

#### Lemma

For any 
$$0<\epsilon<1$$
, and  $k\geq 1$ ,  $\mathsf{Pr}[|p-p^*|\geq\epsilon]\leq 2e^{-rac{k\epsilon^2}{3}}$  .

### Proof.

Recall:  $X_i$  is **1** is  $i^{th}$  ball is black, else **0**.  $B = \sum_{i=1}^{k} X_i$ .  $E[X_i] = p^*$ ,  $E[B] = kp^*$ . p = B/k.  $\Pr[|p - p^*| \ge \epsilon] = \Pr[|\frac{B}{k} - p^*| \ge \epsilon]$  $= \Pr[|B - kp^*| > k\epsilon]$  $= \Pr\left[|B - kp^*| \ge \left(\frac{\epsilon}{p^*}\right)kp^*\right]$ (Chernoff)  $\leq 2e^{-\frac{\epsilon^2}{3p^{*2}}kp^*} = 2e^{-\frac{k\epsilon^2}{3p^*}}$  $(p^* < 1) < 2e^{-\frac{k\epsilon^2}{3}}$ 

2

# Example Summary

The problem was to estimate the fraction of black balls  $p^*$  in a bin filled with white and black balls. Our estimate was  $p = \frac{B}{k}$  instead, where out of k draws (with replacement) B balls turns out black.

#### Markov's Inequality

For any  $k \geq 1$ ,  $\Pr[p \geq 2p^*] \leq \frac{1}{2}$ 

### Chebyshev's Inequality

For any  $0 < \epsilon < 1$ , and  $k \ge 1$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 1/k\epsilon^2$ .

### Chernoff Bound

For any  $0 < \epsilon < 1$ , and  $k \ge 1$ ,  $\Pr[|p - p^*| \ge \epsilon] \le 2e^{-\frac{k\epsilon^2}{3}}$ .

# Part IV

# Randomized QuickSort (Contd.)

# Randomized QuickSort: Recall

### Input: Array A of n numbers. Output: Numbers in sorted order.

# Randomized QuickSort

- **1** Pick a pivot element *uniformly at random* from **A**.
- Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- In the subarrays, and concatenate them.

**Note:** On *every* input randomized **QuickSort** takes  $O(n \log n)$  time in expectation. On *every* input it may take  $\Omega(n^2)$  time with some small probability.

# Randomized QuickSort: Recall

#### Input: Array A of n numbers. Output: Numbers in sorted order.

# Randomized QuickSort

- **1** Pick a pivot element *uniformly at random* from **A**.
- Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- Secursively sort the subarrays, and concatenate them.

**Note:** On *every* input randomized **QuickSort** takes  $O(n \log n)$  time in expectation. On *every* input it may take  $\Omega(n^2)$  time with some small probability.

**Question:** With what probability it takes  $O(n \log n)$  time?

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - 1/n^3$ .

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - 1/n^3$ .

# If n = 100 then this gives $\Pr[Q(A) \le 32n \ln n] \ge 0.999999$ .

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - \frac{1}{n^3}$ .

### Outline of the proof

• k: depth of the recursion. Then  $Q(A) \leq kn$ .

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - \frac{1}{n^3}$ .

- k: depth of the recursion. Then  $Q(A) \leq kn$ .
- Prove that  $k \leq 32 \ln n$  with high probability. Which will imply the result.

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - \frac{1}{n^3}$ .

- k: depth of the recursion. Then  $Q(A) \leq kn$ .
- Prove that  $k \leq 32 \ln n$  with high probability. Which will imply the result.
  - Focus on a single element. Prove that it "participates" in  $> 32 \ln n$  levels with probability at most  $1/n^4$ .
  - By union bound, any of the *n* elements participates in
     > 32 ln *n* levels with probability at most

### Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - \frac{1}{n^3}$ .

- k: depth of the recursion. Then  $Q(A) \leq kn$ .
- Prove that  $k \leq 32 \ln n$  with high probability. Which will imply the result.
  - Focus on a single element. Prove that it "participates" in  $> 32 \ln n$  levels with probability at most  $1/n^4$ .
  - Sy union bound, any of the *n* elements participates in > 32 ln *n* levels with probability at most  $1/n^3$ .

# Informal Statement

Random variable Q(A) = # comparisons done by the algorithm.

We will show that  $\Pr[Q(A) \leq 32n \ln n] \geq 1 - \frac{1}{n^3}$ .

- k: depth of the recursion. Then  $Q(A) \leq kn$ .
- Prove that  $k \leq 32 \ln n$  with high probability. Which will imply the result.
  - Focus on a single element. Prove that it "participates" in  $> 32 \ln n$  levels with probability at most  $1/n^4$ .
  - Sy union bound, any of the *n* elements participates in > 32 ln *n* levels with probability at most  $1/n^3$ .
  - 3 Therefore, all elements participate in  $\leq 32 \ln n$  w.p.  $(1 1/n^3)$ .

• If *k* levels of recursion then *kn* comparisons.

- If k levels of recursion then kn comparisons.
- Fix an element  $s \in A$ . We will track it at each level.
- Let  $S_i$  be the partition containing s at  $i^{th}$  level.
- $S_1 = A$  and  $S_k = \{s\}$ .

# Randomized QuickSort: High Probability Analysis

- If k levels of recursion then kn comparisons.
- Fix an element  $s \in A$ . We will track it at each level.
- Let  $S_i$  be the partition containing s at  $i^{th}$  level.
- $S_1 = A$  and  $S_k = \{s\}$ .
- We call s lucky in  $i^{th}$  iteration, if balanced split:  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ .

# Randomized QuickSort: High Probability Analysis

- If k levels of recursion then kn comparisons.
- Fix an element  $s \in A$ . We will track it at each level.
- Let  $S_i$  be the partition containing s at  $i^{th}$  level.
- $S_1 = A$  and  $S_k = \{s\}$ .
- We call s lucky in  $i^{th}$  iteration, if balanced split:  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ .
- If  $\rho = \#$  lucky rounds in first k rounds, then  $|S_k| \leq (3/4)^{\rho} n$ .

# Randomized QuickSort: High Probability Analysis

- If *k* levels of recursion then *kn* comparisons.
- Fix an element  $s \in A$ . We will track it at each level.
- Let  $S_i$  be the partition containing s at  $i^{th}$  level.
- $S_1 = A$  and  $S_k = \{s\}$ .
- We call s lucky in  $i^{th}$  iteration, if balanced split:  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ .
- If  $\rho = \#$  lucky rounds in first k rounds, then  $|S_k| \leq (3/4)^{\rho} n$ .
- For  $|S_k| = 1$ ,  $\rho = \log_{4/3} n \le 4 \ln n$  suffices.

s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?
  - Clearly,  $\rho = \sum_{i=1}^{k} X_i$ . Let  $\mu = \mathbf{E}[\rho] = \frac{k}{2}$ .

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?
  - Clearly,  $\rho = \sum_{i=1}^{k} X_i$ . Let  $\mu = \mathbf{E}[\rho] = \frac{k}{2}$ .
  - Set  $k = 32 \ln n$  and  $\delta = \frac{3}{4}$ .  $(1 \delta) = \frac{1}{4}$ .

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?
  - Clearly,  $\rho = \sum_{i=1}^{k} X_i$ . Let  $\mu = \mathbf{E}[\rho] = \frac{k}{2}$ .
  - Set  $k = 32 \ln n$  and  $\delta = \frac{3}{4}$ .  $(1 \delta) = \frac{1}{4}$ .

Probability of NOT getting  $4 \ln n$  lucky rounds out of  $32 \ln n$  rounds

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?
  - Clearly,  $\rho = \sum_{i=1}^{k} X_i$ . Let  $\mu = \mathbf{E}[\rho] = \frac{k}{2}$ .
  - Set  $k = 32 \ln n$  and  $\delta = \frac{3}{4}$ .  $(1 \delta) = \frac{1}{4}$ .

Probability of NOT getting  $4 \ln n$  lucky rounds out of  $32 \ln n$  rounds

$$\begin{aligned} \Pr[\rho \leq 4 \ln n] &= \Pr[\rho \leq \frac{k}{8}] \\ &= \Pr[\rho \leq (1 - \delta)\mu] \end{aligned}$$

- s lucky in round i if  $|S_{i+1}| \leq (3/4)|S_i|$  and  $|S_i \setminus S_{i+1}| \leq (3/4)|S_i|$ 
  - $X_i = 1$  if s is lucky in  $i^{th}$  round.
  - **Observation:**  $X_1, \ldots, X_k$  are independent variables.
  - $\Pr[X_i = 1] = \frac{1}{2}$  Why?
  - Clearly,  $\rho = \sum_{i=1}^{k} X_i$ . Let  $\mu = \mathbf{E}[\rho] = \frac{k}{2}$ .
  - Set  $k = 32 \ln n$  and  $\delta = \frac{3}{4}$ .  $(1 \delta) = \frac{1}{4}$ .

Probability of NOT getting  $4 \ln n$  lucky rounds out of  $32 \ln n$  rounds

$$\begin{aligned} \Pr[\rho \le 4 \ln n] &= \Pr[\rho \le \frac{k}{8}] \\ &= \Pr[\rho \le (1 - \delta)\mu] \\ (Chernoff) &\le 2e^{\frac{-\delta^2 \mu}{2}} = 2e^{-\frac{9k}{64}} \\ &= 2e^{-4.5 \ln n} \le \frac{1}{n^4} \end{aligned}$$

• n input elements. Probability that there is some un-lucky element is at most  $\frac{1}{n^4} * n = \frac{1}{n^3}$ .

- n input elements. Probability that there is some un-lucky element is at most  $\frac{1}{n^4} * n = \frac{1}{n^3}$ .
- **Pr**[depth of recursion in **QuickSort** > 32 ln n]  $\leq \frac{1}{n^3}$ .

- n input elements. Probability that there is some un-lucky element is at most  $\frac{1}{n^4} * n = \frac{1}{n^3}$ .
- **Pr**[depth of recursion in **QuickSort** > 32 ln n]  $\leq \frac{1}{n^3}$ .

#### Theorem

With high probability (i.e.,  $1 - \frac{1}{n^3}$ ) the depth of the recursion of **QuickSort** is  $\leq 32 \ln n$ . Due to *n* comparisons in each level, with high probability, the running time of **QuickSort** is  $O(n \ln n)$ .

- n input elements. Probability that there is some un-lucky element is at most  $\frac{1}{n^4} * n = \frac{1}{n^3}$ .
- **Pr**[depth of recursion in **QuickSort** > 32 ln n]  $\leq \frac{1}{n^3}$ .

#### Theorem

With high probability (i.e.,  $1 - \frac{1}{n^3}$ ) the depth of the recursion of **QuickSort** is  $\leq 32 \ln n$ . Due to *n* comparisons in each level, with high probability, the running time of **QuickSort** is  $O(n \ln n)$ .

Q: How to increase the probability?