

*If you hold a cat by the tail
you learn things you cannot learn any other way.*

— Mark Twain

9 Tail Inequalities

The simple recursive structure of skip lists made it relatively easy to derive an upper bound on the expected *worst-case* search time, by way of a stronger high-probability upper bound on the worst-case search time. We can prove similar results for treaps, but because of the more complex recursive structure, we need slightly more sophisticated probabilistic tools. These tools are usually called *tail inequalities*; intuitively, they bound the probability that a random variable with a bell-shaped distribution takes a value in the *tails* of the distribution, far away from the mean.

9.1 Markov's Inequality

Perhaps the simplest tail inequality was named after the Russian mathematician Andrey Markov; however, in strict accordance with Stigler's Law of Eponymy, it first appeared in the works of Markov's probability teacher, Pafnuty Chebyshev.¹

Markov's Inequality. *Let X be a non-negative integer random variable. For any $t > 0$, we have $\Pr[X \geq t] \leq E[X]/t$.*

Proof: The inequality follows from the definition of expectation by simple algebraic manipulation.

$$\begin{aligned}
 E[X] &= \sum_{k=0}^{\infty} k \cdot \Pr[X = k] && \text{[definition of } E[X]\text{]} \\
 &= \sum_{k=0}^{\infty} \Pr[X \geq k] && \text{[algebra]} \\
 &\geq \sum_{k=0}^{t-1} \Pr[X \geq k] && \text{[since } t < \infty\text{]} \\
 &\geq \sum_{k=0}^{t-1} \Pr[X \geq t] && \text{[since } k < t\text{]} \\
 &= t \cdot \Pr[X \geq t] && \text{[algebra]} \quad \square
 \end{aligned}$$

Unfortunately, the bounds that Markov's inequality implies (at least directly) are often very weak, even useless. (For example, Markov's inequality implies that with high probability, every node in an n -node treap has depth $O(n^2 \log n)$. Well, *duh!*) To get stronger bounds, we need to exploit some additional structure in our random variables.

¹The closely related tail bound traditionally called Chebyshev's inequality was actually discovered by the French statistician Irénée-Jules Bienaymé, a friend and colleague of Chebyshev's.

9.2 Sums of Indicator Variables

Recall that random variables X_1, X_2, \dots, X_n are *mutually independent* if and only if

$$\Pr \left[\bigwedge_{i=1}^n (X_i = x_i) \right] = \prod_{i=1}^n \Pr[X_i = x_i]$$

for all possible values x_1, x_2, \dots, x_n . For examples, different flips of the same fair coin are mutually independent, but the number of heads and the number of tails in a sequence of n coin flips are not independent (since they must add to n). Mutual independence of the X_i 's implies that the expectation of the product of the X_i 's is equal to the product of the expectations:

$$\mathbb{E} \left[\prod_{i=1}^n X_i \right] = \prod_{i=1}^n \mathbb{E}[X_i].$$

Moreover, if X_1, X_2, \dots, X_n are independent, then for any function f , the random variables $f(X_1), f(X_2), \dots, f(X_n)$ are also mutually independent.

Suppose $X = \sum_{i=1}^n X_i$ is the sum of n mutually independent random *indicator* variables X_i . For each i , let $p_i = \Pr[X_i = 1]$, and let $\mu = \mathbb{E}[X] = \sum_i \mathbb{E}[X_i] = \sum_i p_i$.

Chernoff Bound (Upper Tail). $\Pr[X > (1 + \delta)\mu] < \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu$ for any $\delta > 0$.

Proof: The proof is fairly long, but it relies on just a few basic components: a clever substitution, Markov's inequality, the independence of the X_i 's, The World's Most Useful Inequality $e^x > 1 + x$, a tiny bit of calculus, and lots of high-school algebra.

We start by introducing a variable t , whose role will become clear shortly.

$$\Pr[X > (1 + \delta)\mu] = \Pr[e^{tX} > e^{t(1+\delta)\mu}]$$

To cut down on the superscripts, I'll usually write $\exp(x)$ instead of e^x in the rest of the proof. Now apply Markov's inequality to the right side of this equation:

$$\Pr[X > (1 + \delta)\mu] < \frac{\mathbb{E}[\exp(tX)]}{\exp(t(1 + \delta)\mu)}.$$

We can simplify the expectation on the right using the fact that the terms X_i are independent.

$$\mathbb{E}[\exp(tX)] = \mathbb{E} \left[\exp \left(t \sum_i X_i \right) \right] = \mathbb{E} \left[\prod_i \exp(tX_i) \right] = \prod_i \mathbb{E}[\exp(tX_i)]$$

We can bound the individual expectations $\mathbb{E}[e^{tX_i}]$ using The World's Most Useful Inequality:

$$\mathbb{E}[\exp(tX_i)] = p_i e^t + (1 - p_i) = 1 + (e^t - 1)p_i < \exp((e^t - 1)p_i)$$

This inequality gives us a simple upper bound for $\mathbb{E}[e^{tX}]$:

$$\mathbb{E}[\exp(tX)] < \prod_i \exp((e^t - 1)p_i) < \exp \left(\sum_i (e^t - 1)p_i \right) = \exp((e^t - 1)\mu)$$

Substituting this back into our original fraction from Markov's inequality, we obtain

$$\Pr[X > (1 + \delta)\mu] < \frac{\mathbb{E}[\exp(tX)]}{\exp(t(1 + \delta)\mu)} < \frac{\exp((e^t - 1)\mu)}{\exp(t(1 + \delta)\mu)} = (\exp(e^t - 1 - t(1 + \delta)))^\mu$$

Notice that this last inequality holds for *all* possible values of t . To obtain the final tail bound, we will choose t to make this bound as tight as possible. To minimize $e^t - 1 - t - t\delta$, we take its derivative with respect to t and set it to zero:

$$\frac{d}{dt}(e^t - 1 - t(1 + \delta)) = e^t - 1 - \delta = 0.$$

(And you thought calculus would never be useful!) This equation has just one solution $t = \ln(1 + \delta)$. Plugging this back into our bound gives us

$$\Pr[X > (1 + \delta)\mu] < (\exp(\delta - (1 + \delta)\ln(1 + \delta)))^\mu = \left(\frac{e^\delta}{(1 + \delta)^{1 + \delta}}\right)^\mu$$

And we're done! □

This form of the Chernoff bound can be a bit clumsy to use. A more complicated argument gives us the bound

$$\Pr[X > (1 + \delta)\mu] < e^{-\mu\delta^2/3} \text{ for any } 0 < \delta < 1.$$

A similar argument gives us an inequality bounding the probability that X is significantly *smaller* than its expected value:

Chernoff Bound (Lower Tail). $\Pr[X < (1 - \delta)\mu] < \left(\frac{e^{-\delta}}{(1 - \delta)^{1 - \delta}}\right)^\mu < e^{-\mu\delta^2/2}$ for any $\delta > 0$.

9.3 Back to Treaps

In our analysis of randomized treaps, we defined the indicator variable A_k^i to have the value 1 if and only if the node with the i th smallest key ('node i ') was a proper ancestor of the node with the k th smallest key ('node k '). We argued that

$$\Pr[A_k^i = 1] = \frac{[i \neq k]}{|k - i| + 1},$$

and from this we concluded that the expected depth of node k is

$$\mathbb{E}[\text{depth}(k)] = \sum_{i=1}^n \Pr[A_k^i = 1] = H_k + H_{n-k} - 2 < 2 \ln n.$$

To prove a worst-case expected bound on the depth of the tree, we need to argue that the *maximum* depth of any node is small. Chernoff bounds make this argument easy, once we establish that the relevant indicator variables are mutually independent.

Lemma 1. For any index k , the $k-1$ random variables A_k^i with $i < k$ are mutually independent. Similarly, for any index k , the $n-k$ random variables A_k^i with $i > k$ are mutually independent.

Proof: To simplify the notation, we explicitly consider only the case $k = 1$, although the argument generalizes easily to other values of k . Fix $n - 1$ arbitrary indicator values x_2, x_3, \dots, x_n . We prove the lemma by induction on n , with the vacuous base case $n = 1$. The definition of conditional probability gives us

$$\begin{aligned} \Pr \left[\bigwedge_{i=2}^n (A_1^i = x_i) \right] &= \Pr \left[\bigwedge_{i=2}^{n-1} (A_1^i = x_i) \wedge A_1^n = x_n \right] \\ &= \Pr \left[\bigwedge_{i=2}^{n-1} (A_1^i = x_i) \mid A_1^n = x_n \right] \cdot \Pr [A_1^n = x_n] \end{aligned}$$

Now recall that $A_1^n = 1$ if and only if node n has the smallest priority, and the other $n - 2$ indicator variables A_1^i depend only on the order of the priorities of nodes 1 through $n - 1$. There are exactly $(n - 1)!$ permutations of the n priorities in which the n th priority is smallest, and each of these permutations is equally likely. Thus,

$$\Pr \left[\bigwedge_{i=2}^{n-1} (A_1^i = x_i) \mid A_1^n = x_n \right] = \Pr \left[\bigwedge_{i=2}^{n-1} (A_1^i = x_i) \right]$$

The inductive hypothesis implies that the variables A_1^2, \dots, A_1^{n-1} are mutually independent, so

$$\Pr \left[\bigwedge_{i=2}^{n-1} (A_1^i = x_i) \right] = \prod_{i=2}^{n-1} \Pr [A_1^i = x_i].$$

We conclude that

$$\Pr \left[\bigwedge_{i=2}^n (A_1^i = x_i) \right] = \Pr [A_1^n = x_n] \cdot \prod_{i=2}^{n-1} \Pr [A_1^i = x_i] = \prod_{i=1}^{n-1} \Pr [A_1^i = x_i],$$

or in other words, that the indicator variables are mutually independent. \square

Theorem 2. *The depth of a randomized treap with n nodes is $O(\log n)$ with high probability.*

Proof: First let's bound the probability that the depth of node k is at most $8 \ln n$. There's nothing special about the constant 8 here; I'm being generous to make the analysis easier.

The depth is a sum of n indicator variables A_k^i , as i ranges from 1 to n . Our Observation allows us to partition these variables into two mutually independent subsets. Let $d_<(k) = \sum_{i < k} A_k^i$ and $d_>(k) = \sum_{i > k} A_k^i$, so that $\text{depth}(k) = d_<(k) + d_>(k)$. If $\text{depth}(k) > 8 \ln n$, then either $d_<(k) > 4 \ln n$ or $d_>(k) > 4 \ln n$.

Chernoff's inequality, with $\mu = \mathbb{E}[d_<(k)] = H_k - 1 < \ln n$ and $\delta = 3$, bounds the probability that $d_<(k) > 4 \ln n$ as follows.

$$\Pr[d_<(k) > 4 \ln n] < \Pr[d_<(k) > 4\mu] < \left(\frac{e^3}{4^4} \right)^\mu < \left(\frac{e^3}{4^4} \right)^{\ln n} = n^{\ln(e^3/4^4)} = n^{3-4 \ln 4} < \frac{1}{n^2}.$$

(The last step uses the fact that $4 \ln 4 \approx 5.54518 > 5$.) The same analysis implies that $\Pr[d_>(k) > 4 \ln n] < 1/n^2$. These inequalities imply the crude bound $\Pr[\text{depth}(k) > 4 \ln n] < 2/n^2$.

Now consider the probability that the treap has depth greater than $10 \ln n$. Even though the distributions of different nodes' depths are *not* independent, we can conservatively bound the probability of failure as follows:

$$\Pr \left[\max_k \text{depth}(k) > 8 \ln n \right] = \Pr \left[\bigwedge_{k=1}^n (\text{depth}(k) > 8 \ln n) \right] \leq \sum_{k=1}^n \Pr[\text{depth}(k) > 8 \ln n] < \frac{2}{n}.$$

This argument implies more generally that for any constant c , the depth of the treap is greater than $c \ln n$ with probability at most $2/n^{c \ln c - c}$. We can make the failure probability an arbitrarily small polynomial by choosing c appropriately. \square

This lemma implies that any search, insertion, deletion, or merge operation on an n -node treap requires $O(\log n)$ time with high probability. In particular, the expected *worst-case* time for each of these operations is $O(\log n)$.

Exercises

1. Prove that for any integer k such that $1 < k < n$, the $n - 1$ indicator variables A_k^i with $i \neq k$ are *not* mutually independent. [Hint: Consider the case $n = 3$.]
2. Recall from Exercise 1 in the previous note that the expected number of descendants of any node in a treap is $O(\log n)$. Why doesn't the Chernoff-bound argument for depth imply that, with high probability, *every* node in a treap has $O(\log n)$ descendants? The conclusion is clearly bogus—Every treap has a node with n descendants!—but what's the hole in the argument?
3. A *heater* is a sort of dual treap, in which the priorities of the nodes are given, but their search keys are generate independently and uniformly from the unit interval $[0, 1]$. You can assume all priorities and keys are distinct.
 - (a) Prove that for any r , the node with the r th smallest *priority* has expected depth $O(\log r)$.
 - (b) Prove that an n -node heater has depth $O(\log n)$ with high probability.
 - (c) Describe algorithms to perform the operations INSERT and DELETEMIN in a heater. What are the expected worst-case running times of your algorithms? In particular, can you express the expected running time of INSERT in terms of the priority rank of the newly inserted item?